

## RESEARCH ARTICLE

# Aggregation of incomplete preference rankings: Robustness analysis of the $ZM_{II}$ -technique

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## Abstract

A common group decision-making problem is that in which: (a) several *judges* express their *subjective* preference rankings regarding some *objects* of interest and (b) these rankings should then be aggregated into a *collective* judgement. The authors recently developed an aggregation technique – denominated “ $ZM_{II}$ ” – aggregating these rankings into a *ratio* scaling of the objects, which represents the solution to the decision-making problem of interest. This technique also includes a flexible *response mode*, which tolerates *incomplete* rankings and can, therefore, be adapted to various practical contexts, such as quality improvement activities, field surveys, product-comparison surveys, etc.

The aim of this article is proposing an original approach to verify the robustness of the  $ZM_{II}$ -technique under the influence of various factors, especially those concerned with the degree of “completeness” of preference rankings (e.g., number of objects identified by judges, whether these objects are ordered or not, etc.). The methodology in use relies on the simulation of several thousand decision-making problems, in order to organically study the effect of the factors of interest. Results show a certain robustness of the  $ZM_{II}$ -technique, even under relatively “unfavourable” practical conditions, characterized by very incomplete preference rankings. Description is supported by instructive examples.

## KEYWORDS

degree of agreement, degree of completeness, factorial analysis, *generalized least squares*, incomplete preference ranking, quality improvement, robustness analysis, Thurstone's *Law of Comparative Judgement*,  $ZM_{II}$ -technique

## 1 | INTRODUCTION

A general group-decision problem – which is transversal to many scientific disciplines, such as design/manufacturing decision making, quality improvement, quality management, etc. – is structured as follows (Das & Mukherjee, 2007; Franceschini & Maisano, 2019a; Keeney & Raiffa, 1993; Wang, Huang, Yan, & Du, 2017):

- some *objects* ( $o_1, o_2, \dots$ ) should be evaluated and compared on the basis of a certain *attribute*;

- some *judges* ( $j_1, j_2, \dots$ ) make their *subjective* judgements on these objects. In addition, judges may refrain from evaluating a subset of the objects because of practical impediment or lack of adequate knowledge;
- The ultimate goal is to aggregate the judges' judgements into a single *collective judgement*.

Numerous aggregation techniques have been developed so far, which differ in: (a) the *response mode* through which subjective judgements are expressed (e.g., ratings, rankings, paired-comparison relationships, etc.), (b) the type of *aggregation model* (e.g., heuristic,

statistical, fuzzy models, etc.), and (c) the nature of the *collective judgement* (e.g., object rankings, ordinal/interval/ratio scale values, etc.). For an exhaustive discussion of the existing techniques, we refer the reader to (Arrow, 1963; DeVellis, 2016; Fishburn, 1989; Franceschini, Galetto, Maisano, & Mastrogiacono, 2017; Nurmi, 1983).

A key element for the success of a generic aggregation technique is the simplicity of the response mode (Franceschini, Galetto, & Maisano, 2019; Harzing et al., 2009; Paruolo, Saisana, & Saltelli, 2013), for example, some authors showed that comparative judgements of objects (e.g., “ $o_i$  is more/less preferred than  $o_j$ ”) are generally simpler and more reliable than judgements in absolute terms (e.g., “the degree of the attribute of  $o_i$  is low/intermediate/high”) (Edwards, 1957; Harzing et al., 2009).

The authors have recently developed an aggregation technique – denominated “ $ZM_{II}$ ” – which relies on the postulates and simplifying assumptions of the Thurstone’s *Law of Comparative Judgement* (LCJ) (Edwards, 1957; Franceschini & Maisano, 2018; Franceschini & Maisano, 2020a; Thurstone, 1927) and embodies the *Generalized Least Squares* (GLS) method (Kariya & Kurata, 2004; Ross, 2014). This technique includes a relatively versatile response mode that tolerates “incomplete” rankings, that is, rankings only including the objects at the top/bottom (Franceschini & Maisano, 2019a). For the sake of simplicity, a decision-making problem characterized by this type of rankings will be hereafter referred to as “incomplete ranking problem” or, even more simply, as “incomplete problem.” On the other hand, a decision-making problem characterized by “complete” rankings will be hereafter referred to as “complete ranking problem” or, even more simply, as “complete problem.”

The flexible response mode of the  $ZM_{II}$ -technique makes it adaptable to a variety of practical contexts, where judges do not have the concentration to formulate complete rankings; for example, problems with a relatively large number of objects, field surveys, product-comparison surveys, customer-satisfaction survey, etc. (Chen & Cheng, 2010; Harzing et al., 2009; Lagerspetz, 2016).

In addition, the  $ZM_{II}$ -technique allows (a) to construct a *ratio* scaling of the objects, which represents the output solution of the decision-making problem of interest, and (b) to estimate the uncertainty of this resulting scaling, by “propagating” the uncertainty of input data (Franceschini & Maisano, 2019a; Roberts, 1979; Zhang, Guo, & Chen, 2016).

An important requirement of the  $ZM_{II}$ -technique is that, apart from the objects to be evaluated – that is,  $o_1, o_2, \dots, o_n$ , which will be

hereafter classified as “regular” objects – preference rankings also include two fictitious “dummy” objects, that is,  $o_z$  and  $o_m$ , which will be described in the next section.<sup>1</sup>

This paper aims at proposing an original approach to organically verify the robustness of the  $ZM_{II}$ -technique, depending on the “degree of completeness”<sup>2</sup> and other characteristic factors of the (incomplete) decision-making problems. By the term “robustness,” it is meant the ability of the  $ZM_{II}$ -technique to provide a collective judgement comparable to that which would be obtained if the judges formulated complete rankings. This investigation is an essential step to substantiate the practical convenience and effectiveness of the  $ZM_{II}$ -technique in real-world practical contexts.

For the purpose of example, let us consider the two decisional problems in Table 1, both characterized by four judges ( $j_1$  to  $j_4$ ) formulating their individual rankings of four objects ( $o_1$  to  $o_4$ ). Objects are in this case alternative design concepts of some pocket projectors, which have to be ranked in terms of *ease of use* (i.e., attribute of interest).

In the first case, rankings are *complete* while in the second case they are *incomplete*, reflecting an “unfavourable” practical context, in which judges do not have the possibility to formulate complete rankings. It can be seen that the information content of incomplete rankings represents a subset of the information content of complete rankings. For example,  $o_1$  is in the top and  $o_4$  is in the bottom of the incomplete ranking by  $j_1$ , but nothing is known about the mutual relationships between the intermediate objects; on the other hand, in the corresponding complete ranking, this information is given:  $o_2 \sim o_3$ .

Considering that the  $ZM_{II}$  aggregation technique can be applied to both the decision-making problems, a question may arise: “*To what extent will the solution of an incomplete problem be distorted (due to the lower information content), compared to that of a corresponding complete problem?*” Or, reversing the question: “*To what extent does the  $ZM_{II}$ -technique tolerate incomplete problems, with no significant distortion of the solution compared to corresponding complete problems?*”

From a methodological point of view, a large number of decision-making problems with different “structural” factors (e.g., number of objects, number of judges, etc.) will be simulated; then, the solution of complete problems will be compared with the solutions of a number of corresponding incomplete ones (e.g., problems characterized by rankings with the more/less preferred objects only and/or without the dummy objects, etc.). The solution of a generic complete problem

(1) Complete problem	(2) incomplete problem	
Rankings	Rankings	Comment
$j_1: o_1 > (o_2 \sim o_3) > o_4$	$o_1 > \dots > o_4$	Only the object at the top and the one at the bottom of the ranking are identified.
$j_2: o_1 > o_2 > (o_3 \sim o_4)$	$o_1 > o_2 > \dots$	Only the first two objects at the top of the ranking are identified.
$j_3: o_2 > (o_1 \sim o_3 \sim o_4)$	$o_2 > \dots$	Only the first object at the top of the ranking is identified.
$j_4: o_1 \sim o_2 \sim o_3 \sim o_4$	$o_1 \sim o_2 \sim o_3$	Object $o_4$ is not assessed by the judge since it is not well-known.

**TABLE 1** Example of two decisional problems (one complete and one incomplete)

will be used as a “gold standard” to check the goodness of the solutions of the corresponding incomplete problems. Reversing the perspective, the present study will tell us how far we can go in making preference rankings more and more incomplete – without deteriorating the solution excessively.

The rest of the paper is organized into five sections. The section “Background information” briefly recalls the response mode and the basic principles of the  $ZM_{II}$ -technique. The section “Methodology” illustrates the methodology to examine the robustness of the solution of an incomplete problem, with respect to that of a “source” complete problem; special attention will be devoted to the description of a set of (simulated) factorial experiments. The “Results” section presents and discusses the results of these experiments in detail, highlighting their practical consequences; the description is supported by an extensive use of explanatory graphs. The “Concluding remarks” section summarizes the original contributions of this paper and its practical implications, limitations, and suggestions for future research. Further details on the results of this research are contained in the Appendix A section.

## 2 | BACKGROUND INFORMATION

This section briefly recalls the  $ZM_{II}$ -technique and provides preparatory information to understand the rest of the paper. The presentation is organized into three subsections concerning: (a) the *response mode*, (b) the *artificial deterioration of complete rankings* (into incomplete ones), and (c) the *rationale of the aggregation technique* in use.

### 2.1 | Response mode

A prerequisite of the  $ZM_{II}$ -technique is that each judge involved in the problem formulates a preference ranking of the objects – that is, a sequence of objects in order of preference, with the more preferred ones in the top positions and the less preferred ones in the bottom ones.

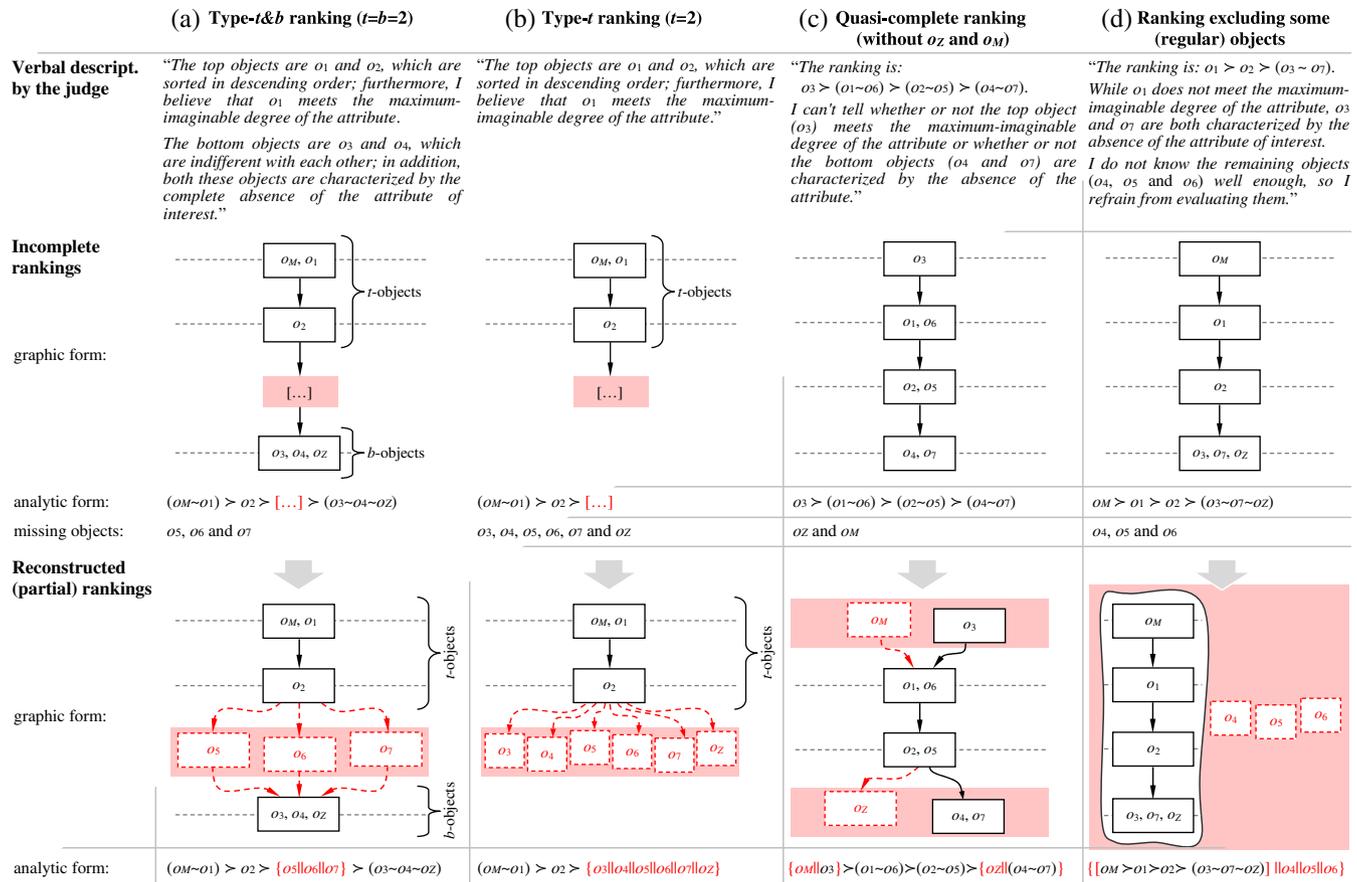
Apart from *regular* objects ( $o_1, o_2, \dots$ ), judges may also include two (fictitious) *dummy* objects in their rankings: that is, one ( $o_Z$ ) corresponding to the *absence* of the attribute of interest, and one ( $o_M$ ) corresponding to the *maximum-imaginable* degree of the attribute (Franceschini & Maisano, 2020a). When dealing with these dummy objects, two important requirements should be considered:

1.  $o_Z$  should be positioned at the bottom of a preference ranking, that is, there should not be any other object with preference lower than  $o_Z$ . In the case the attribute of another object is judged to be absent, that object will be considered indifferent to  $o_Z$  and positioned at the same hierarchical level.
2.  $o_M$  should be positioned at the top of a preference ranking, that is, there should not be any other object with preference higher than  $o_M$ . In the case the attribute of another object is judged to be the maximum-imaginable, that object will be considered indifferent to  $o_M$  and positioned at the same hierarchical level.

In the best cases, judges formulate complete preference rankings. Borrowing the language from Order Theory, any ranking including all (regular and dummy) objects, according to a hierarchical sequence with relationships of *strict preference* and/or *indifference* only, can be classified as *linear* (Gierz, Hofmann, Keimel, Mislove, & Scott, 2003). Unfortunately, the formulation of complete preference rankings may sometimes be problematic (Harzing et al., 2009). To overcome this obstacle, a flexible response mode that tolerates incomplete preference rankings can be adopted. Below is a list of some possible types of incomplete preference rankings.

- Preference rankings including only the more preferred objects (or “ $t$ -objects,” where “ $t$ ” stands for “top”) and the less preferred ones (or “ $b$ -objects,” where “ $b$ ” stands for “bottom”); these rankings will be hereafter denominated “Type- $t&b$ .” The  $t$  parameter is conventionally defined as the number of regular objects (i.e., excluding the two dummy objects) within the  $t$ -objects, while the  $b$  parameter is conventionally defined as the number of regular objects within  $b$ -objects. In the example in Figure 1a, the judge merely specifies the two objects at the top and the two objects at the bottom of the ranking, therefore  $t = b = 2$ .
- Preference rankings including only the more preferred objects (i.e.,  $t$ -objects) among those available; see the example in Figure 1b, in which  $t = 2$ . From now on, these rankings will be denominated “Type- $t$ .”
- Preference rankings not including the two dummy objects ( $o_Z$  and  $o_M$ ), for example, in the case judges find it difficult to envisage them. These preference rankings will be classified as “quasi-complete” if they include all regular objects; see the example in Figure 1c.
- Combining the previous three types of incomplete preference rankings, one can obtain Type- $t&b$  or Type- $t$  preference rankings that do not include the dummy objects.
- To contemplate the fact that judges may not be able to evaluate certain objects – for example, since they are not familiar with them – preference rankings excluding some objects will also be tolerated (see Figure 1d).

The upper part of Figure 1 exemplifies some judges' verbal descriptions from which the previous types of incomplete rankings can be deduced. Figure 1 also shows that a generic incomplete ranking can be transformed into a “reconstructed” ranking, which includes all the (dummy and regular) objects, with the addition of appropriate incomparability relationships. Borrowing the language from Order Theory, these other rankings can be classified as *partial*, that is, apart from *strict preference* (e.g., “ $o_i > o_j$ ”) and *indifference* (e.g., “ $o_i \sim o_j$ ”) relationships, they may also contain *incomparability* relationships (e.g., “ $o_i \parallel o_j$ ”) among the objects (Gierz et al., 2003). For example, considering Type- $t&b$  rankings, the objects that are not considered by judges can be allocated at an intermediate hierarchical level with respect to the  $t$ - and  $b$ -objects, with mutual *incomparability* relationships. As for Type- $t$  rankings, the objects that are not considered by judges can be allocated at a lower hierarchical level with respect to the  $t$ -objects. As for



**FIGURE 1** Example of four different types of incomplete rankings, including seven regular objects (*o<sub>1</sub>* to *o<sub>7</sub>*) and two dummy objects (*o<sub>Z</sub>* and *o<sub>M</sub>*). These rankings can be turned into reconstructed (partial) rankings, which include all the (regular and dummy) objects; for ease of understanding, the reconstructed parts are marked in red colour

the rankings that do not include *o<sub>Z</sub>* and *o<sub>M</sub>*, they can be reconstructed in compliance with the following constraints:

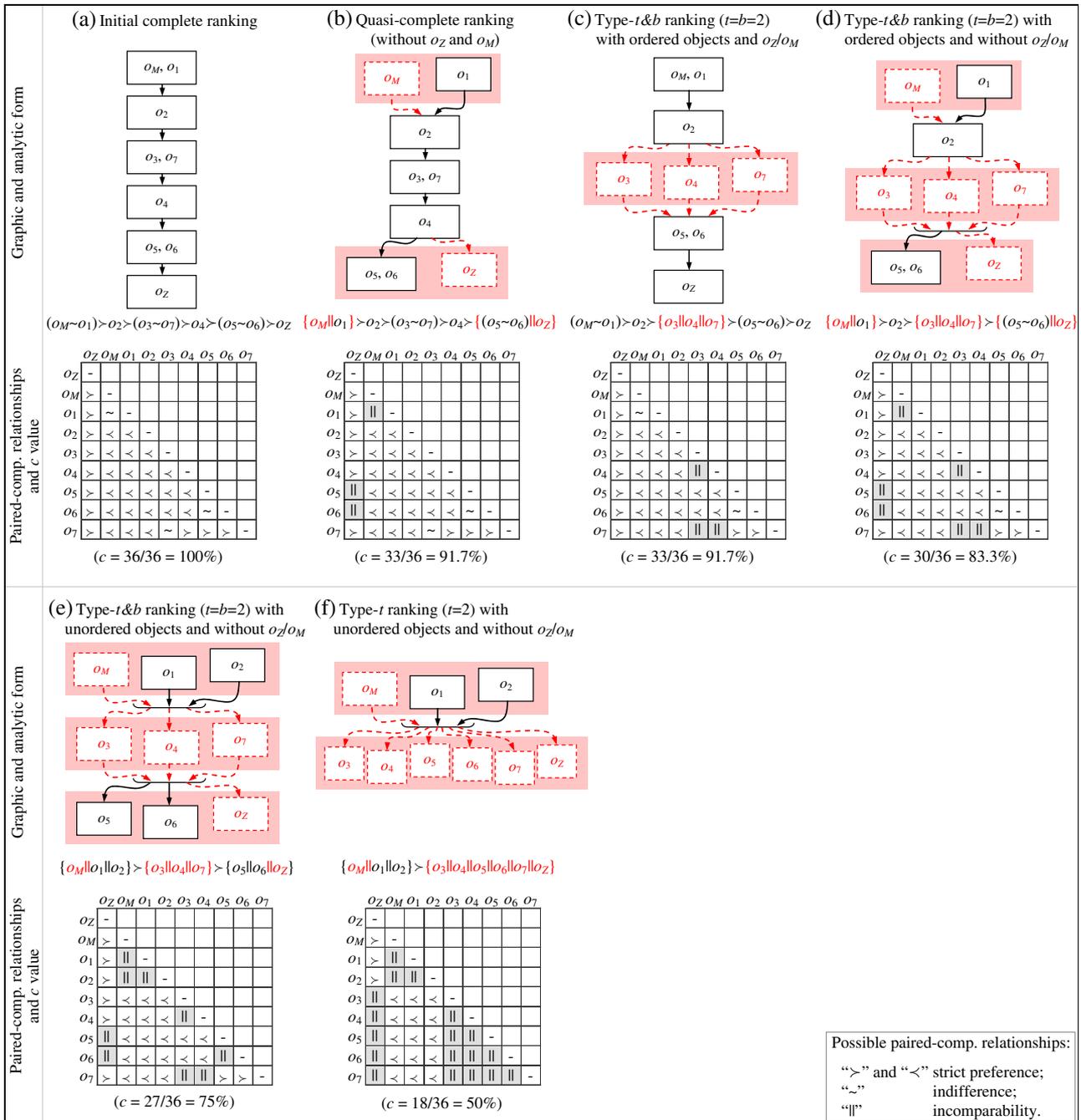
1. the dummy object *o<sub>M</sub>* should – by definition – be positioned at the same hierarchical level of the top object(s) or above. To take this hesitation into account, a relationship of incomparability between *o<sub>M</sub>* and the top object(s) can be introduced. For example, when reconstructing the ranking in Figure 2b, *o<sub>M</sub>* is inserted at the top of the ranking, with a relationship of incomparability with respect to the top object, *o<sub>1</sub>*;
2. the dummy object *o<sub>Z</sub>* should – by definition – be positioned at the same hierarchical level of the bottom object(s) or below. To take this other hesitation into account, a relationship of incomparability between *o<sub>Z</sub>* and the bottom object(s) can be introduced. For example, when reconstructing the ranking in Figure 2b, *o<sub>Z</sub>* is inserted at the bottom of the ranking, with a relationship of incomparability with respect to the bottom objects, *o<sub>5</sub>* and *o<sub>6</sub>*.

Let us now make a brief digression on the meaning of the term “incomparability.” In the multiple criteria decision analysis (MCDA) field, it generally depicts a comparison between two objects, in which

a judge opts neither for a strict preference nor indifference relationship (Bana e Costa & Vansnick, 1999). This hesitation is generally due to lacking and/or contradictory information available, concerning the two objects. Table 2 summarizes a plurality of possible practical situations that lead to incomparability.

Firstly, we notice that there is no incomparability if-and-only-if the available information meets the following two requisites at the same time: it should be (a) in the appropriate *amount* and (b) *non-contradictory* (see the practical situations at points 1 and 2 of Table 2). Since the above requisites are not met in the remaining eight practical situations (at points 3–10), they all result in incomparability.

For the sake of simplicity, the present study will be limited to situations of incomparability that can be ascribed to that at point 6: “*o<sub>i</sub>* could be preferred or equivalent to *o<sub>j</sub>*, but the lack of relevant information precludes determining the most appropriate relationship.” The incomparability of two objects will, therefore, be seen as a sort of hesitation between the relationships of strict preference and indifference, excluding conflicting situations. Of course, the authors are aware that this meaning of the term “incomparability” is narrower than that in other MCDA contexts.



**FIGURE 2** Example of gradual deterioration of a (complete) preference ranking (a), generating several incomplete rankings: (b)–(f); for ease of understanding, the “reconstructed” parts are marked in red colour

## 2.2 | Artificial deterioration of complete rankings

Let us now make a brief digression to show a contrivance that will be used later in our analysis: that is, the artificial “deterioration” of a complete ranking to generate a set of incomplete preference rankings that are compatible<sup>3</sup> with it. Precisely, this contrivance will be used to generate incomplete rankings, artificially reproducing practical circumstances where the formulation of complete ones can be problematic.

Let us focus on the example in Figure 2, in which an initial complete preference ranking is given; Figure 2a shows the decomposition

of this ranking into paired-comparison relationships of strict preference and indifference. Next, the initial complete ranking can be gradually deteriorated and turned into several incomplete preference rankings; for example, consider the quasi-complete ranking in Figure 2b, the Type- $t$ & $b$  ranking in Figure 2c, the Type- $t$ & $b$  ranking in Figure 2d, etc. It can be noticed that for these incomplete rankings, new paired-comparison relationships of incomparability gradually replace those of strict preference and indifference in the complete ranking.

The *compatibility* between the initial complete ranking and the respective incomplete rankings is given by the fact that – excluding

**TABLE 2** Possible situations occurring when comparing two generic objects ( $o_i$  and  $o_j$ ), depending on the *amount* and the *congruence* of the information available

Practical situation	Information available		Relation
1. $o_i$ is preferred to $o_j$ as the attribute of $o_i$ is judged definitely better than that of $o_j$ , from any point of view.	Proper amount	Congruent	$o_i > o_j$
2. $o_i$ and $o_j$ are indifferent because their attributes are indifferent from any point of view.	Proper amount	Congruent	$o_i \sim o_j$
3. $o_i$ and $o_j$ are in a conflicting position due to the clearly contradictory information about them.	Proper amount	Contradictory	$o_i \parallel o_j$
4. $o_i$ and $o_j$ could be either indifferent or conflicting, but the lack of relevant information in any of the two directions leads to hesitation.	Lacking	?	$o_i \parallel o_j$
5. $o_i$ and $o_j$ could be either indifferent or conflicting, but excessive and contradictory information in any of the two directions leads to hesitation.	Excessive amount	Contradictory	$o_i \parallel o_j$
6. $o_i$ could be preferred or equivalent to $o_j$ , but the lack of relevant information precludes determining the most appropriate relationship.	Lacking	Congruent	$o_i \parallel o_j$
7. $o_i$ could be preferred or equivalent to $o_j$ , but excessive and contradictory information precludes determining the most appropriate relationship.	Excessive amount	Contradictory	$o_i \parallel o_j$
8. $o_j$ cannot be preferred to $o_i$ , but the lack of relevant information precludes determining whether (a) $o_i$ is preferred to $o_j$ or (b) they are conflicting.	Lacking	?	$o_i \parallel o_j$
9. $o_j$ cannot be preferred to $o_i$ , but excessive and contradictory information precludes determining whether (a) $o_i$ is preferred to $o_j$ or (b) they are conflicting.	Excessive amount	Contradictory	$o_i \parallel o_j$
10. $o_i$ could be preferred to $o_j$ , but bad information leads to hesitation.	Lacking	Contradictory	$o_i \parallel o_j$

Note: This scheme was constructed by reworking the content of Tsoukiàs and Vincke (1995).

the paired-comparison relationships of incomparability – the remaining ones are identical. In general, a generic complete ranking can be deteriorated in different ways, generating a large set of incomplete rankings that are compatible with it.

Returning to the artificially generated incomplete rankings, the *degree of completeness* of a generic  $k$ -th ranking can be quantitatively described by the synthetic indicator:

$$c = \frac{\text{No. of "usable" paired comparison relations in the preference ranking}}{\binom{n}{2}} \quad (1)$$

which expresses the fraction of “usable” paired-comparison relationships – that is, of strict preference or indifference – with respect to the total ones:  $\binom{n}{2} = n \cdot (n-1)/2$ , where  $n$  is the total number of objects of the problem; the adjective “usable” indicates that these are the only relationships that contribute to the solution of the decision-making problem of interest. By way of example, we determined the  $c$  values related to the rankings exemplified in Figure 2 (below the tables containing the paired-comparison relationships). This indicator tends to increase while rankings become more and more complete; for complete preference rankings,  $c$  is obviously 1.

Interestingly, even rankings that are apparently very incomplete may contain a relevant portion of usable paired-comparison relationships. For example, consider the Type- $t$  ranking in Figure 2f, in which only the two more preferred regular objects are selected but not ordered; despite the apparently high degree of incompleteness, half of the usable paired-comparison relationships are still preserved ( $c = 50\%$ ).

The indicator  $c$  can be extended from a single preference ranking to sets of  $m$  preference rankings – such as those characterizing a decision-making problem with  $m$  judges. We thus define a new aggregated indicator ( $\bar{c}$ ), which depicts the overall degree of completeness:

$$\bar{c} = \frac{\sum_{k=1}^m \text{No. of "usable" paired comparison relations in the } k^{\text{th}} \text{ preference ranking}}{\sum_{k=1}^m \text{Total no. of paired comparison relations in the } k^{\text{th}} \text{ preference ranking}} \\ = \frac{\sum_{k=1}^m c_k \cdot \binom{n}{2}}{m \cdot \binom{n}{2}} = \frac{\sum_{k=1}^m c_k}{m} \quad (2)$$

$c_k$  is the degree of completeness of a generic  $k$ -th ranking.

Equation (2) also shows that  $\bar{c}$  can also be interpreted as the arithmetic mean of the  $c$  values related to the set of preference rankings under consideration.

For more examples about the artificial deterioration of complete rankings into incomplete ones, we refer the reader to the section "Example of generation of incomplete rankings" (in the Appendix A).

### 2.3 | Rationale of the aggregation technique

The mathematical formalization of the problem relies on the postulates and simplifying assumptions of the LCJ by Thurstone (1927), who postulated the existence of a *psychological/psychophysical continuum*, in which objects are positioned depending on the degree of a certain attribute. The position of a generic  $i$ -th object ( $f_i$ ) is postulated to be distributed normally, in order to reflect the intrinsic judge-to-judge variability:  $f_i \sim N(x_i, \sigma_{x_i}^2)$ , where  $x_i$  and  $\sigma_{x_i}^2$  are the unknown mean value and variance related to the degree of the attribute of that object. In addition, the distributions of different objects are considered equally dispersed and equally correlated with each other (Edwards, 1957; Thurstone, 1927). Considering two generic objects,  $o_i$  and  $o_j$ , and having introduced further simplifying hypotheses (Edwards, 1957; Thurstone, 1927), it can be asserted that:

$$p_{ij} = P[(f_i - f_j) > 0] = 1 - \Phi[-(x_i - x_j)], \quad (3)$$

which expresses the probability ( $p_{ij}$ ) that the position of  $f_i$  is higher than that of  $f_j$ ,  $\Phi$  is the cumulative distribution function of the standard normal distribution  $z \sim N(0, 1)$ . Although  $p_{ij}$  is unknown, it can be estimated using the information contained in a set of judgements expressed by a number ( $m$ ) of judges (Edwards, 1957; Thurstone, 1927). For more information on the estimation of the  $p_{ij}$  values, based on the positioning of the objects in the (reconstructed) rankings of judges, see Franceschini and Maisano (2019a).

From Equation (3) it can be inferred that:

$$x_i - x_j = -\Phi^{-1}(1 - p_{ij}). \quad (4)$$

Extending the reasoning to all possible pairs of objects, among which relationships of strict preference or indifference can be deduced (Franceschini & Maisano, 2019a), and introducing further simplifying assumptions, an over-determined system of ( $q$ ) equations [like Equation (4)] can be then obtained. Since this system consists of linear equations with respect to the unknowns (i.e.,  $x_i$  values), it can be expressed in matrix form as:

$$\begin{pmatrix} \vdots \\ \sum_{k=1}^n (a_{hk} \cdot x_k) - b_h = 0 \quad \forall h \in [0, q] \\ \vdots \end{pmatrix} \Rightarrow \mathbf{A} \cdot \mathbf{X} - \mathbf{B} = \mathbf{0} \quad (5)$$

$\mathbf{X} = [\dots, x_i, \dots]^T \in \mathbb{R}^{n \times 1}$  is the column vector containing the unknowns of the problem,  $a_{hk}$  is a generic element of matrix  $\mathbf{A} \in \mathbb{R}^{q \times n}$ ,

and  $b_h$  is a generic element of vector  $\mathbf{B} \in \mathbb{R}^{q \times 1}$ . For details on the construction of  $\mathbf{A}$  and  $\mathbf{B}$ , see (Gulliksen, 1956).

Then, this system can be solved by applying the GLS method (Kariya & Kurata, 2004), which allows to obtain an estimate of the mean degree of the attribute of each object:

$$\mathbf{X} = (\mathbf{A}^T \cdot \mathbf{W} \cdot \mathbf{A})^{-1} \cdot \mathbf{A}^T \cdot \mathbf{W} \cdot \mathbf{B}, \quad (6)$$

where  $\mathbf{W}$  is a (squared) matrix encapsulating the uncertainty related to the equations of the system; a practical way to define  $\mathbf{W}$  is to apply the *Multivariate Law of Propagation of Uncertainty* (MLPU) to the system in Equation (6), referring to the input variables affected by uncertainty (Kariya & Kurata, 2004), that is, the  $p_{ij}$  values; for details, see Franceschini and Maisano (2019a).

The scale values in  $\mathbf{X}$  are expressed on an arbitrary *interval* scale (Thurstone, 1927). The uncertainty of the solution can be estimated through a covariance matrix  $\Sigma_X$ , which can be obtained by propagating the uncertainty of input data (i.e.,  $p_{ij}$  values), through the following relationship:

$$\Sigma_X = (\mathbf{A}^T \cdot \mathbf{W} \cdot \mathbf{A})^{-1}. \quad (7)$$

Through the following transformation, the scale value of a generic  $i$ -th object ( $x_i$ ) is transformed into a new scale value ( $y_i$ ), which is defined in the conventional range  $[0, 10]$ :

$$\mathbf{Y} = \mathbf{Y}(\mathbf{X}) = [\dots, y_i(\mathbf{X}), \dots]^T = \left[ \dots, 100 \cdot \frac{x_i - x_Z}{x_M - x_Z}, \dots \right]^T, \quad (8)$$

where  $x_Z$  and  $x_M$  are the scale values of  $o_Z$  and  $o_M$ , respectively, in the initial interval scale;  $x_i$  is the scale value of a generic  $i$ -th object in the initial interval scale;  $y_i$  is the scale value of a generic  $i$ -th object in the new scale (Franceschini & Maisano, 2019a). Since scale  $y$  "inherits" the *interval* property from scale  $x$  and has a conventional zero point that corresponds to the absence of the attribute (i.e.,  $y_Z = 0$ ), it can be reasonably considered as a *ratio* scale, without any conceptually prohibited "promotion" (Franceschini et al., 2019).

Next, the uncertainty related to the elements in  $\mathbf{Y} = [\dots, y_i, \dots]^T \in \mathbb{R}^{n \times 1}$  can be determined by applying a classic approach borrowed from Metrology to Equation (8): the so-called *Delta Method*, also referred as *Law of Propagation of Uncertainty* or *Error Transmission Formula* (JCGM 100:2008, 2008). It is thus obtained:

$$\Sigma_Y = \mathbf{J}_{Y(\mathbf{X})} \cdot \Sigma_X \cdot \mathbf{J}_{Y(\mathbf{X})}^T, \quad (9)$$

where  $\mathbf{J}_{Y(\mathbf{X})} \in \mathbb{R}^{n \times n}$  is a Jacobian matrix containing the partial derivatives related to the equations of the system in Equation (8), with respect to the elements of  $\mathbf{X}$ ; for details, see (Franceschini & Maisano, 2019a).

Combining Equations (7) and (9),  $\Sigma_Y$  can be expressed as:

$$\Sigma_Y = J_{Y(X)} \cdot \left[ (A^T \cdot W \cdot A)^{-1} \right] \cdot J_{Y(X)}^T \quad (10)$$

Assuming that the  $p_{ij}$  and  $y_i$  values are normally distributed, a 95% confidence interval related to each  $y_i$  value can be computed as:

$$y_i \pm U_i = y_i \pm 2 \cdot \sigma_i \quad \forall i \quad (11)$$

$U_i$  is the so-called *expanded uncertainty* of  $y_i$  with a coverage factor  $k = 2$  and  $\sigma_i = \sqrt{\Sigma_{Y, (ij)}}$  (JCGM 100:2008, 2008).

### 3 | METHODOLOGY

#### 3.1 | General approach and response indicators

With the aim of investigating the robustness of the solution provided by the  $ZM_{II}$ -technique for incomplete problems, the methodological approach is articulated into several general points (see the flowchart in Figure 3):

- Numerous complete problems are randomly generated, determining the relevant solutions.
- These complete problems are then artificially deteriorated into incomplete ones, determining the new corresponding solutions (see the example in Figure 2). It is assumed that each complete problem has the same number of incomplete problems (i.e., 19, as explained in detail below).
- The solution of each incomplete problem is compared with the solution of the corresponding (source) complete problem, which can be interpreted as a sort of “gold standard.” In fact, it can be demonstrated that the  $ZM_{II}$ 's solution to a complete problem coincides with that one provided by Thurstone's LCJ (Arbuckle &

Nugent, 1973; Edwards, 1957; Franceschini & Maisano, 2019a; Gulliksen, 1956; Thurstone, 1927), which is a very consolidated aggregation technique, only applicable to complete problems. This is a sort of guarantee of *plausibility* of the  $ZM_{II}$ 's results.

The robustness analysis is performed using two appropriate response indicators. The first one ( $\bar{\varepsilon}$ ) expresses the deviation of the solution of a certain incomplete problem from that of the corresponding (source) complete problem, and it is structured as a *Root-Mean-Square Error* (RMSE) of the deviations between the  $y_i$  values resulting from these two problems (Ross, 2014):

$$\bar{\varepsilon} = \sqrt{\frac{\sum_{i=1}^n (y_i - y_{i(\text{complete})})^2}{n}} \quad (12)$$

$y_i$  is the scale value of the  $i$ -th object, resulting from the solution of the incomplete problem (i.e.,  $Y = [\dots, y_i, \dots]^T$ );  $y_{i(\text{complete})}$  is the scale value of the  $i$ -th object, resulting from the solution of the complete problem (i.e.,  $Y_{(\text{complete})} = [\dots, y_{i(\text{complete})}, \dots]^T$ );

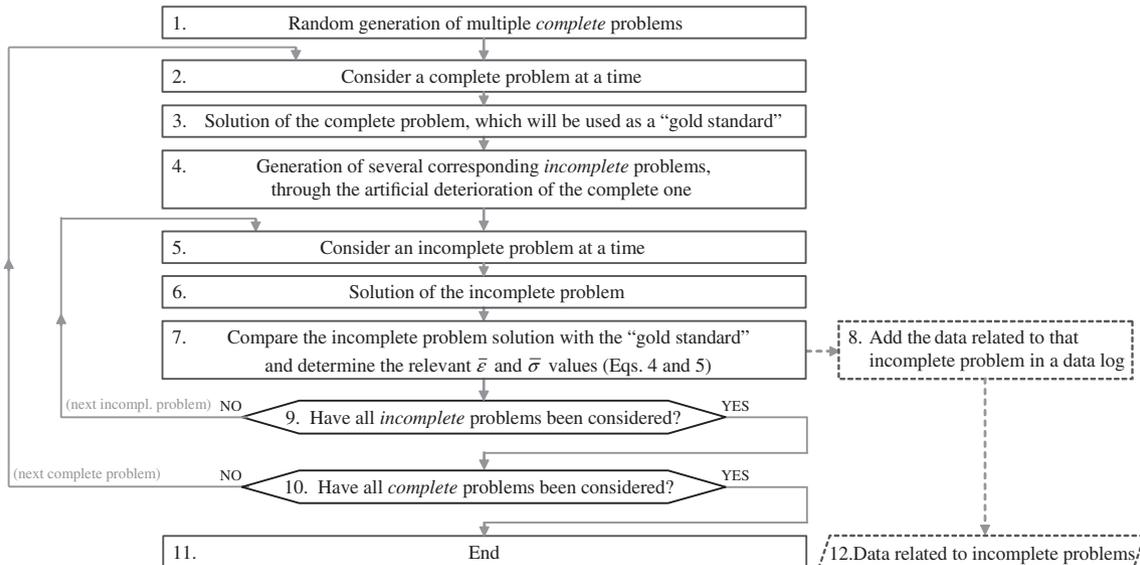
$n$  is the number of (dummy and regular) objects.

The closer the  $y_{i(\text{complete})}$  values get to the  $y_i$  values, the more  $\bar{\varepsilon}$  will tend to decrease; in this sense, this indicator is a measure of (*in*)*accuracy* (JCGM 200:2012, 2012). The calculation of this response indicator for complete rankings will obviously “degenerate” into  $\bar{\varepsilon} = 0$ .

The second response indicator is:

$$\bar{\sigma} = \sqrt{\frac{\sum_{i=1}^n \sigma_i^2}{n}} \quad (13)$$

$\sigma_i^2$  is the variance related to the scale value ( $y_i$ ) of the  $i$ -th object, that is, one of the elements contained in the diagonal of the covariance



**FIGURE 3** Flowchart representing the proposed methodological approach.  $\bar{\varepsilon}$  and  $\bar{\sigma}$  are two response indicators that will be defined later on

matrix of  $Y$  (i.e.,  $\Sigma_Y$ , defined in Equations (9) and (10)). This response indicator – which can be calculated for both complete and incomplete problems – depicts the average *dispersion* of the solution.

### 3.2 | Factorial simulations

This section describes the structured factorial (simulated) experiments that have been carried out. The scheme in Figure 4 summarizes the (sub-)factors considered, dividing them into three families:

1. *Structural factors*, which characterize a complete problem and the corresponding incomplete problems derived from it;
2. *Deterioration sub-factors*, which determine the way a complete problem is artificially “deteriorated,” originating several incomplete problems.
3. *Completeness factor*, which quantifies the degree of completeness of each problem.

Several complete problems are initially generated, according to the logic illustrated in the following three points.

1. The number of *regular* objects ( $n_{reg}$ , which does not take into account the two dummy objects<sup>4</sup>) is varied over nine levels:  $n_{reg} = 4, 5, 6, 7, 8, 9, 10, 11, \text{ and } 12$ .
  2. For each of the above  $n_{reg}$  values, the number of judges ( $m$ ) – which obviously coincides with the number of preference rankings – is varied over six levels:  $m = 5, 10, 15, 20, 25, \text{ and } 30$ .
  3. For each combination of  $n_{reg}$  and  $m$  values, four complete decision-making problems are randomly generated; these problems will be associated with one category expressing the qualitative degree of agreement of judges: *very low*, *low*, *intermediate*, and *high*. Below is a detailed description of the procedure for the random generation of the complete problems and their classification into the above four categories.
- a. A random scaling of the regular objects is generated, assigning a (random) scale value between 0 and 100 to each object.

- b. The scale value of each object is randomly distorted by introducing a uniformly distributed error in the range  $\pm \Delta$ ; the value of  $\Delta$  is conventionally set to 10, 20, 30, and 40, respectively for a *very low*, *low*, *intermediate*, and *high* degree of inter-judge agreement. The resulting scale value is then rounded to the nearest 10 (e.g., 20, 30, 40, etc.). Of course, this scale value can be outside the range  $[0, 100]$ , with no effect on the subsequent steps.
- c. The latter scaling is then translated into a new complete preference ranking.
- d. Steps (b) and (c) are repeated  $m$  times, obtaining a complete problem with  $m$  randomly generated complete preference rankings.

The inter-judge degree of agreement of a resulting complete problem can be better quantified through the so-called Kendall's *coefficient of concordance* (Franceschini & Maisano, 2019b; Kendall, 1962; Legendre, 2010):

$$W = \frac{12 \cdot \sum_{i=1}^n R_i^2 - 3 \cdot m^2 \cdot n \cdot (n+1)^2}{m^2 \cdot n \cdot (n^2 - 1) - m \cdot \sum_{j=1}^m T_j}, \quad (14)$$

where

- $n$  is the total number of (dummy and regular) objects;
- $R_i$  is the sum of the rank positions for the  $i$ -th object ( $o_i$ ), that is,  $R_i = \sum_{j=1}^m r_{ij}$ , in which terms  $r_{ij}$  represent the rank of  $o_i$  according to the  $j$ -th judge;
- $n$  is the total number of objects;
- $m$  is the total number of rankings;
- $T_j$  is a correction factor for ties,<sup>5</sup>  $T_j = \sum_{i=1}^{g_j} (t_i^3 - t_i)$ , in which  $t_j$  is the number of tied ranks in the  $i$ -th group of tied ranks (where a group is a set of values having a constant tied rank) and  $g_j$  is the number of groups of ties in the set of ranks (ranging from 1 to  $n$ ) for judge  $j$ . Thus,  $T_j$  is the correction factor required for the set of ranks for judge  $j$ . Note that if there are no tied ranks for judge  $j$ ,  $T_j = 0$ .

The range of  $W$  is between 0 (full disagreement) and 1 (full agreement). For more information on the construction of  $W$ , see

**FIGURE 4** Scheme of the (sub) factors characterizing the proposed factorial simulations. The dashed arrow indicates that the degree of completeness of a problem ( $\bar{c}$ ) is influenced by the four deterioration sub-factors

(a) *Structural factors*:

- Number of regular objects ( $n_{reg}$ );
- Number of judges ( $m$ );
- Qualitative degree of agreement among judges;
- Kendall's concordance coefficient ( $W$ ).

(b) *Deterioration sub-factors*:

- Type of rankings (hereafter abbreviated as “Ranking type”);
- Ability of the judge to manage  $o_Z$  and  $o_M$  (hereafter abbreviated as “Manage  $o_Z/o_M$ ?”);
- Value of  $t$  and/or  $b$  (hereafter abbreviated as “ $t/b$  value”);
- Ability of the judge to order the  $t$ - and/or  $b$ -objects (hereafter abbreviated as “Order  $t/b$ -objects?”).

(c) *Completeness factor*:

- Overall degree of completeness of the problem ( $\bar{c}$ ).

(Franceschini & Maisano, 2019b; Kendall, 1962; Legendre, 2010). Apparently, the  $W$  value associated with a generic complete problem is likely to be correlated to the corresponding qualitative degree of inter-judge agreement; this will be empirically demonstrated in the next section.

Thus,  $9 \cdot 6 \cdot 4 = 216$  (i.e., number of  $n_{reg}$  levels times number of  $m$  levels times number of levels of the qualitative inter-judge degree of agreement) complete decision-making problems were generated. For each of them, 19 incomplete problems are then generated by changing the deterioration sub-factors, according to the contrivance anticipated in the sub-section "Artificial deterioration of complete rankings." Precisely:

1. A single (quasi-complete) problem with  $m$  quasi-complete rankings. These incomplete rankings realistically represent practical situations in which the judges, while patiently considering all the regular objects, do find it difficult to manage  $o_Z$  and  $o_M$ .
2. Three incomplete problems with  $m$  Type- $t$ & $b$  rankings including dummy objects, where  $t$ - and  $b$ -objects are ordered. For the first, second, and third of these problems,  $t$  and  $b$  were set to 1, 2, and 3 respectively; in other words, apart from the dummy objects, the incomplete rankings include only 1, 2, and 3 more and less preferred regular objects.<sup>6</sup> These rankings can be appropriate when judges are not required to include all the regular objects, for example, due to lack of time, concentration, etc..
3. Three incomplete problems characterized by  $m$  Type- $t$ & $b$  rankings without dummy objects, where  $t$ - and  $b$ -objects are ordered. For the first, second, and third of these problems,  $t$  and  $b$  were set to 1, 2, and 3 respectively. These rankings may be appropriate when judges are not required to include all the regular objects and find it difficult to manage  $o_Z$  and  $o_M$ .
4. Three incomplete problems characterized by  $m$  Type- $t$ & $b$  rankings without dummy objects and with *unordered*  $t$ - and  $b$ -objects. For the first, second, and third of these problems,  $t$  and  $b$  were set to 1, 2, and 3 respectively. The degree of completeness of these problems is significantly lower than that of the incomplete problems at point (c), since judges only select the  $t$ - and  $b$ -objects, without ordering them.
5. Three incomplete problems with  $m$  Type- $t$  rankings including dummy objects, where  $t$ - and  $b$ -objects are ordered. For the first, second, and third of these problems,  $t$  was set to 1, 2, and 3 respectively. These rankings can be appropriate for decision-making problems aimed at ranking the more preferred objects only, neglecting the less preferred ones.
6. Three incomplete problems characterized by  $m$  Type- $t$  rankings without dummy objects, where  $t$ -objects are ordered. For the first, second, and third of these problems,  $t$  was set to 1, 2, and 3 respectively. These rankings may be appropriate for decision-making problems aimed at ranking the more preferred objects only, in situations in which judges find it difficult to manage  $o_Z$  and  $o_M$ .
7. Three incomplete problems characterized by  $m$  Type- $t$  rankings without dummy objects and with *unordered*  $t$ -objects. For the first,

second, and third of these problems,  $t$  was set to 1, 2, and 3 respectively.

The section "Example of generation of incomplete rankings" (in the Appendix A) further exemplifies the generation of incomplete rankings. As described before, a number of incomplete problems with variable degrees of incompleteness can be generated; for example, the "least incomplete" ones are the quasi-complete problems, while the most severe incompleteness is the one related to type- $t$  problems with  $t = 1$ . The degree of completeness of the resulting incomplete problems will be qualitatively estimated using the indicator  $\bar{c}$  [in Equation (2)].

The solution of each of the above incomplete problems will be compared with that of the corresponding (source) complete problem, determining the two response indicators,  $\bar{e}$  and  $\bar{o}$  [in Equations (12) and (13)]. The total number of (complete and incomplete) decision-making problems generated will therefore be:  $(9 \cdot 6 \cdot 4) \cdot (1 + 19) = 4,320$  (see also Figure 5).

The overall degree of completeness of a specific incomplete problem will be evaluated through  $\bar{c}$  [see Equation (2)]. Regarding the inter-judge degree of agreement, Equation (14) cannot be applied directly to incomplete problems; in fact,  $W$  is only applicable to complete problems. For simplicity, it was assumed that incomplete problems "inherit" the  $W$  values of the corresponding (source) complete problems. The same can be extended to the qualitative degree of agreement (*very low*, *low*, *intermediate*, and *high*). We plan to overcome this approximation in the future by developing a revised version of  $W$ , which can also be applied to incomplete rankings (Franceschini & Maisano, 2020b).

## 4 | RESULTS

A spreadsheet (which is available in the additional material) reports the detailed results of the 4,320 simulated decision-making problems. The following three subsections illustrate, respectively, (a) a qualitative and (b) a quantitative analysis of the above results, and (c) a regressive model reproducing the effects of the predominant factors.

### 4.1 | Quantitative analysis

Figures 6 and 7 contain the *main effects plots*,<sup>7</sup> representing the effect of the major examined factors ( $n_{reg}$ ,  $m$ ,  $\bar{c}$ , and  $W$ ) on the two responses, that is,  $\bar{e}$  and  $\bar{o}$  respectively. For practical reasons,  $W$  was used as a quantitative indicator of the degree of inter-judge agreement. The box-plot in Figure 8 shows the relatively strong link between the  $W$  value of a generic complete problem and the respective qualitative degree of inter-judge agreement.

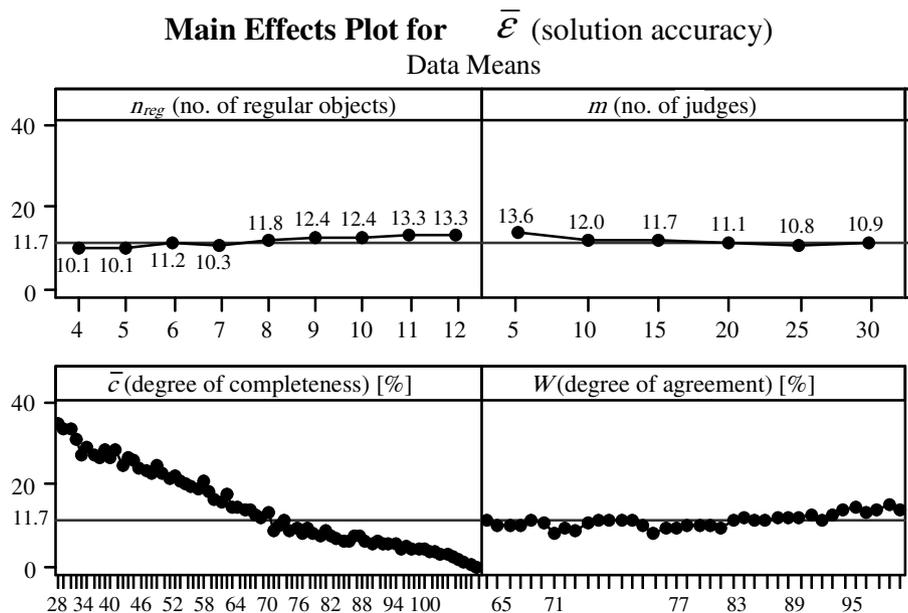
In addition, the degree of completeness of a specific problem is assessed – at least in the first instance – using only  $\bar{c}$  and neglecting the relevant deterioration sub-factors (which are specified in Figure 4b).

1. Change the number of regular objects ( $n_{reg}$ ) at nine levels (i.e., from 4 to 12).	(x9)
2. Change the number of judges ( $m$ ) at six levels (i.e., 5, 10, 15, 20, 25 and 30).	(x6)
3. Change the degree of agreement of judges at four levels ( <i>very low</i> , <i>low</i> , <i>intermediate</i> and <i>high</i> ).	(x4)
4.1 Generate a complete problem with rankings matching the factors set at points 1, 2 and 3. Solve the problem; the resulting solution will be considered as a “gold standard”.	(1)
4.2 Deterioration of the afore-mentioned complete problem into the following (19) incomplete problems:	
4.2.1 Obtain a quasi-complete problem (i.e., with regular objects only).	(1)
4.2.2 Change the $t$ and $b$ parameters at three levels (i.e., 1, 2 and 3).	
4.2.2.1 Obtain a problem consisting of type- $t$ & $b$ rankings with ordered $t$ - and $b$ -objects, and with $o_Z$ and $o_M$ .	(1x3)
4.2.2.2 Obtain a problem consisting of type- $t$ & $b$ rankings, with ordered $t$ - and $b$ -objects, and without $o_Z$ and $o_M$ .	(1x3)
4.2.2.3 Obtain a problem consisting of type- $t$ & $b$ rankings, with unordered $t$ - and $b$ -objects.	(1x3)
4.2.2.4 Obtain a problem consisting of type- $t$ rankings with ordered $t$ -objects, and with $o_Z$ and $o_M$ .	(1x3)
4.2.2.5 Obtain a problem consisting of type- $t$ rankings, with ordered $t$ -objects, and without $o_Z$ and $o_M$ .	(1x3)
4.2.2.6 Obtain a problem consisting of type- $t$ rankings, with unordered $t$ -objects.	(1x3)
4.2.3 Solve the problems at points 4.2.1 and 4.2.2, and compare the resulting solutions with the “gold standard”.	

**Total number of simulated problems:  $9 \cdot 6 \cdot 4 \cdot 20 = 4,320$**

**FIGURE 5** Synthetic description of the factorial simulations

**FIGURE 6** Minitab main effect plot of the major examined factors ( $n_{reg}$ ,  $m$ ,  $\bar{c}$ , and  $W$ ) on the first response ( $\bar{\epsilon}$ )

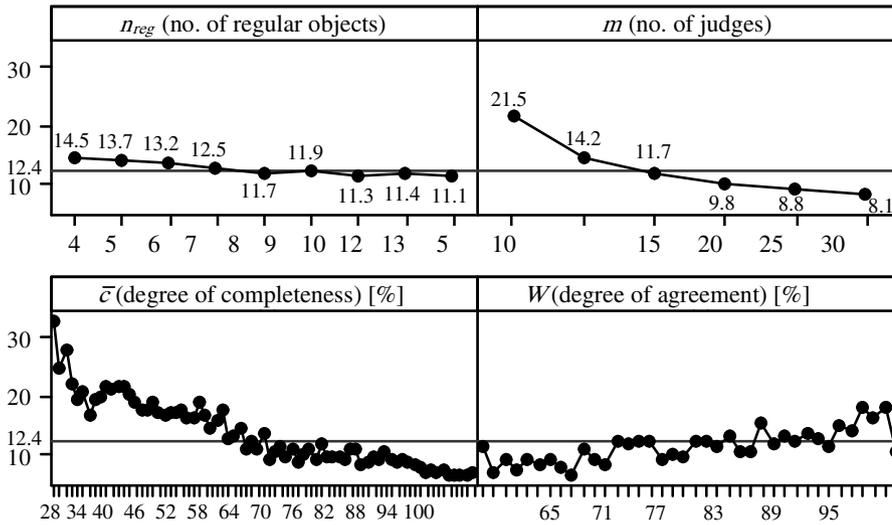


Here are some specific comments on the graphs in Figures 6 and 7.

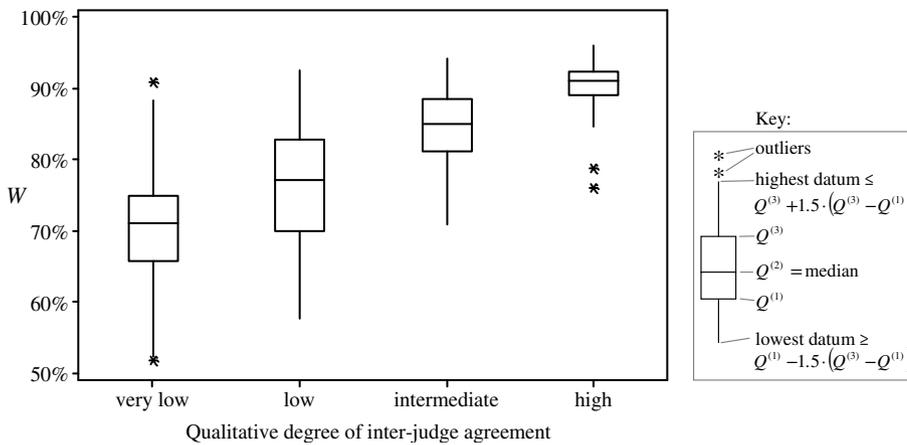
- The factor with the predominant effect on both responses is  $\bar{c}$ , corroborating the hypothesis that the incompleteness of preference rankings contributes to deteriorate the problem solution significantly, both in terms of accuracy ( $\bar{\epsilon}$ ) and dispersion ( $\bar{\sigma}$ ). Nevertheless, the aggregation technique in use proved to be robust, since it allowed to determine solutions that were not exaggeratedly different from the “gold standard,” even for highly incomplete problems (e.g., with  $\bar{c} < 50\%$ ).

- $W$  seems to have a rather weak effect on both responses. Curiously, it seems that a certain disagreement among judges may contribute to reduce both  $\bar{\epsilon}$  and  $\bar{\sigma}$ . In fact, problems with a relatively low  $W$  value result in a more homogeneous distribution of the non-usable relationships of incomparability (among the possible paired comparisons), with a consequent benefit for the solution accuracy and dispersion.
- Factors,  $n_{reg}$  and  $m$ , seem to have not-very-relevant effects on the response,  $\bar{\epsilon}$ . The slight effects in Figure 6 are due to the procedure of random generation of the incomplete problems. In fact, for Type- $t$ & $b$  or Type- $t$  rankings,  $t$  and  $b$  were set to 1, 2, or

**Main Effects Plot for  $\bar{\sigma}$  (solution dispersion)**  
Data Means

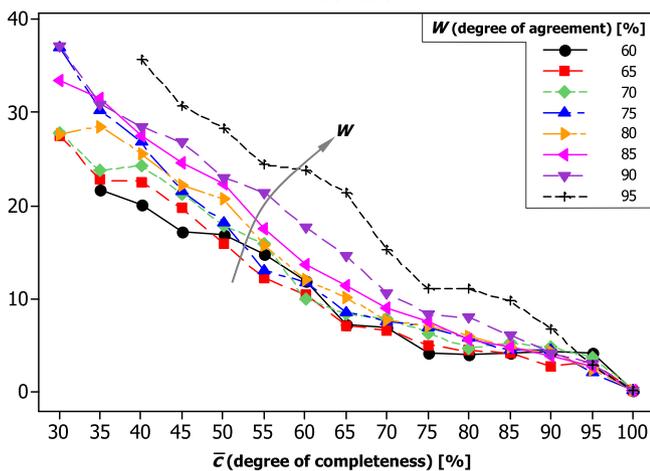


**FIGURE 7** Minitab main effect plot of the major examined factors ( $n_{reg}$ ,  $m$ ,  $\bar{c}$ , and  $W$ ) on the second response ( $\bar{\sigma}$ )



**FIGURE 8** Minitab Box-plot related to the distributions of the  $W$  values, for the complete problems characterized by a certain qualitative degree of agreement (very low, low, intermediate, and high). A relatively strong link between the two indicators can be observed

**$\bar{\epsilon}$  as a function of  $\bar{c}$  and  $W$**   
Data Means



**FIGURE 9** Minitab plot for  $\bar{\epsilon}$  as a function of the two most influential factors,  $\bar{c}$  and  $W$ . The “iso- $W$ ” curves refer to problems with  $W$  values “rounded” to those reported in the legend (i.e., 60, 65, 70%, etc.), considering a resolution of 5%

3, regardless of the total number of regular objects ( $n_{reg}$ ) of the problem. For a certain  $t/b$  value, the degree of completeness of rankings with relatively large  $n_{reg}$  values will reasonably be lower than that of rankings with relatively low  $n_{reg}$  values, with a consequent growth of  $\bar{\epsilon}$ . In addition, as  $m$  grows (more judges)  $W$  will tend to decrease (more probability to obtain discordant rankings), with a consequent decrease in  $\bar{\epsilon}$  (see previous point).

Regarding the response  $\bar{\sigma}$ , the effect of  $n_{reg}$  is irrelevant while that of  $m$  seems relevant. A plausible justification of the latter effect is that the variability of the input tends to decrease while increasing  $m$ , and, therefore, the variability of the solution (depicted by  $\bar{\sigma}$ ) will tend to decrease too (Franceschini & Maisano, 2019a).

#### 4.2 | Qualitative analysis

In order to qualitatively judge the presence of interactions between  $\bar{c}$  and  $W$ , a plot<sup>8</sup> of  $\bar{\epsilon}$  as a function of these factors was constructed (see

**TABLE 3** Pearson correlation table for the main factors ( $n$ ,  $m$ ,  $\bar{c}$ , and  $W$ ) and responses ( $\bar{\varepsilon}$  and  $\bar{\sigma}$ )

Variable	$n_{reg}$	$m$	$\bar{c}$	$W$	$\bar{\varepsilon}$	$\bar{\sigma}$
$n_{reg}$	1					
$m$	0	1				
$\bar{c}$	-0.266*	0.001	1			
$W$	-0.341*	-0.126*	0.084*	1		
$\bar{\varepsilon}$	0.119*	-0.088*	-0.843*	0.125*	1	
$\bar{\sigma}$	-0.145*	-0.561*	-0.580*	0.258*	0.732*	1

\* $p$  value for the hypothesis test of the correlation coefficient being zero is lower than .001.

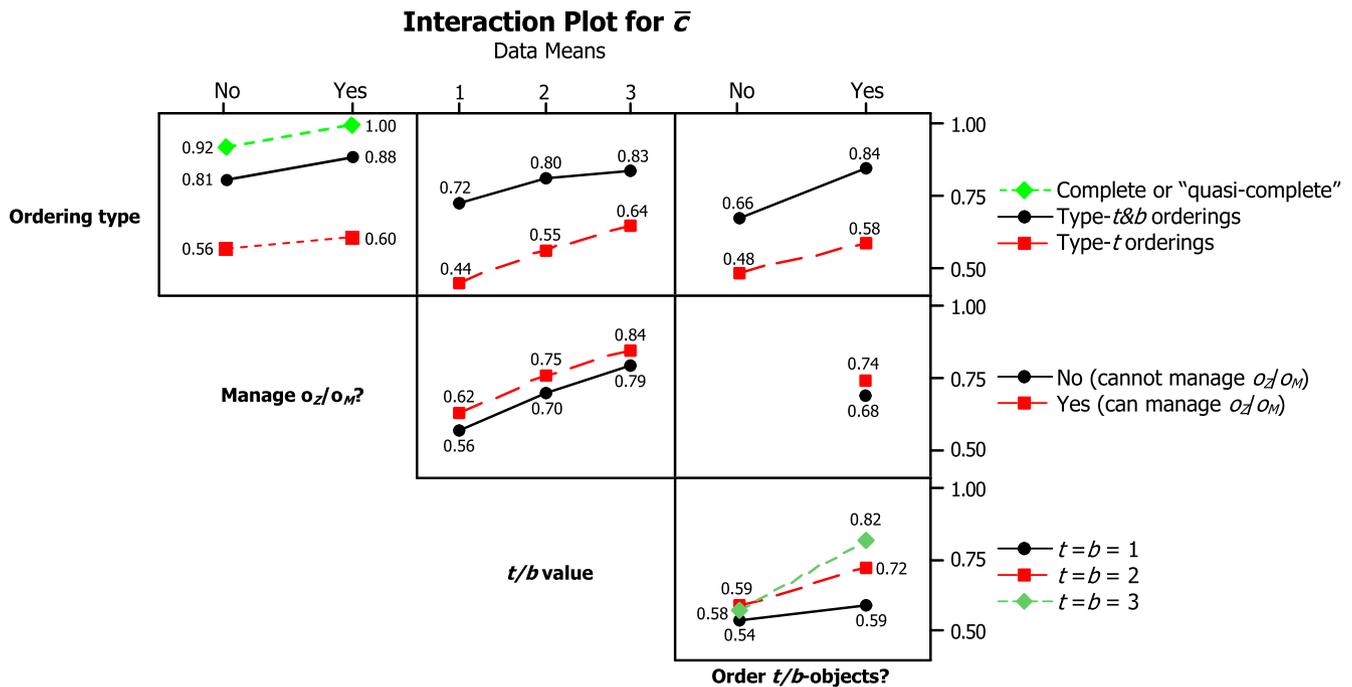
**FIGURE 10** Interaction plot for  $\bar{c}$ , considering the four deterioration sub-factors: "Ranking type," "Manage  $o_z/o_M$ ," "t/b value," and "Order t/b-objects?"

Figure 9). The curves represented in this graph – which can be approximately considered as "iso- $W$ " – are not perfectly "parallel" to each other, denoting a slight correlation between  $\bar{c}$  and  $W$ .

A further quantitative confirmation of the afore-illustrated results is given by Table 3, which contains the Pearson<sup>9</sup> correlation coefficients ( $\rho$ ) between all the possible pairs of variables under consideration (i.e., main factors,  $n_{reg}$ ,  $m$ ,  $\bar{c}$ , and  $W$ , and responses,  $\bar{\varepsilon}$  and  $\bar{\sigma}$ ). For each possible pair, the  $\rho$  value is associated with a corresponding  $p$  value for the hypothesis test of the correlation coefficient being zero (i.e., absence of correlation). Cases in which this value is lower than 0.001 – that is, cases of rejection of the null hypothesis that there is no correlation – are those marked with the symbol "\*."

According to the proposed correlation analysis, the major factor affecting both responses is  $\bar{c}$ , followed by  $W$ . In addition, we note that  $W$  is positively related with  $\bar{\varepsilon}$  and  $\bar{\sigma}$ ; this means that a certain degree of disagreement between the rankings fosters the accuracy and precision of the solution. This behaviour certainly depends on the intrinsic

characteristics of the  $ZM_{II}$  aggregation technique, with particular reference to the propagation of the uncertainty of the input data (Franceschini & Maisano, 2019a).

Excluding the correlations involving  $n_{reg}$  and  $m$ , since they are related to the way incomplete preference rankings are randomly generated,<sup>10</sup> it can be noted that the correlation between  $\bar{c}$  and  $W$  is relatively weak ( $\rho \approx 0.084$ ), confirming the impression gained by analysing Figure 9.

Let us now return to the major factor affecting responses, that is,  $\bar{c}$ , which can be affected by the four deterioration sub-factors: "Ranking type," "Manage  $o_z/o_M$ ," "t/b value," and "Order t/b-objects?." The interaction plot in Figure 10 shows that these sub-factors seem to be uncorrelated with each other. Not surprisingly, their mutual Pearson's correlation coefficients are null. On the other hand, the predominant sub-factor affecting  $\bar{c}$  is the "Ranking type," with  $\rho \approx -0.741$ ; in fact, Figure 10 shows that Type- $t$  rankings tend to make  $\bar{c}$  decrease dramatically, deteriorating the accuracy of the solution. We checked that

the solution accuracy tends to worsen considerably for the less preferred objects, especially due to the relatively lower information content concerning these objects.

The factor  $\bar{c}$  is affected by the sub-factors “t/b value” and “Order t/b-objects?,” with  $\rho$  values of 0.363 and 0.369, respectively. On the other hand, the impact of the sub-factor “Manage  $\sigma_Z/\sigma_M$ ?” is significantly lower ( $\rho \approx 0.154$ ).

### 4.3 | Regression model

To further confirm the above quantitative results, this sub-section illustrates the construction of a regression model, which links the

response,  $\bar{\varepsilon}$ , to the most influential factors ( $\bar{c}$  and  $W$ ). This analysis also provides an estimate of the so-called *effect size* of the factors themselves (Levine & Hullett, 2002).

This section focuses on a regression model to link the response,  $\bar{\varepsilon}$ , with the predominant factors,  $\bar{c}$  and  $W$ , for incomplete decision-making problems. This model enriches the analysis presented in the section “Results.”

The data related to the decision-making problems described in the section “Factorial simulations” have been used to construct the model. Considering Figure 6 – which shows the patterns of  $\bar{\varepsilon}(r)$  and  $\bar{\varepsilon}(W)$  – a second-order polynomial model was chosen. Being quadratic with respect to  $\bar{c}$  and  $W$ , this model seems to well represent the previous graph patterns:

$$\bar{\varepsilon} = K_1 + K_2 \cdot \bar{c} + K_3 \cdot W + K_4 \cdot \bar{c}^2 + K_5 \cdot W^2 + K_6 \cdot \bar{c} \cdot W. \quad (15)$$

#### Best Subsets Regression: $\bar{\varepsilon}$ versus $\bar{c}$ , $W$ , $\bar{c}^2$ , $W^2$ , $\bar{c} \cdot W$

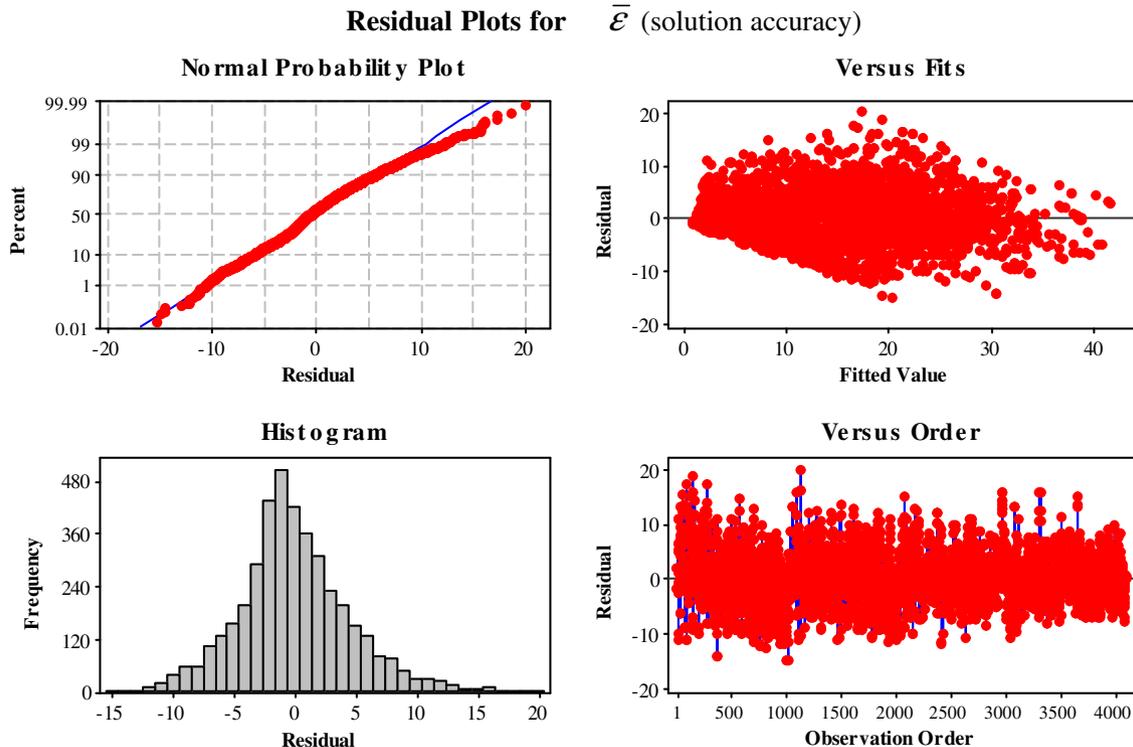
Vars	R-Sq	R-Sq(adj)	Mallows Cp	S	(K <sub>2</sub> ) $\bar{c}$	(K <sub>3</sub> ) $W$	(K <sub>4</sub> ) $\bar{c}^2$	(K <sub>5</sub> ) $W^2$	(K <sub>6</sub> ) $\bar{c} \cdot W$
1	66.9	66.9	1528.3	5.4154	X				
1	62.4	62.4	2283.6	5.7763		X			
2	71.9	71.9	700.4	4.9894	X	X			X
2	71.9	71.9	702.1	4.9903	X		X		X
3	74.8	74.8	223.2	4.7261	X	X	X		
3	74.7	74.7	239.4	4.7353	X	X	X		
4	76.0	76.0	22.7	4.6106	X	X	X	X	X
4	75.9	75.8	53.3	4.6284	X	X	X	X	X
5	76.2	76.1	6.0	4.6003	X	X	X	X	X

**FIGURE 11** Results obtained from Minitab Best-Subsets tool. The above table suggests that the model with the five terms  $\bar{c}$ ,  $W$ ,  $\bar{c}^2$ ,  $W^2$ , and  $\bar{c} \cdot W$  is relatively precise and unbiased because its Mallows' Cp (6.0) is closest to the number of predictors plus the constant (6)

It is important to notice the presence of the last term ( $K_6 \cdot \bar{c} \cdot W$ ), which accounts for the interaction between  $\bar{c}$  and  $W$ .

With the support of the Minitab Best-Subsets tool, it was confirmed the importance of all terms (see results in Figure 11).

Since the variance of the response variable ( $\bar{\varepsilon}$ ) is not homogeneous, a simple linear regression is not perfectly suitable. In particular, heteroscedasticity<sup>11</sup> may have the effect of giving too much weight to data subsets, where the error variance is larger, when estimating coefficients. To reduce the standard error associated with coefficient estimates in regression, in which homoscedasticity is violated, a common approach is to weight observations by the reciprocal of the estimated point variance<sup>12</sup> (Box et al., 1978).



**FIGURE 12** Minitab residual plots resulting from the (weighted) regression analysis

**FIGURE 13** Results of the (weighted) regression analysis**(Weighted) Regression Analysis:  $\bar{\varepsilon}$  versus  $\bar{c}$ ,  $W$ ,  $\bar{c}^2$ ,  $W^2$  and  $\bar{c} \cdot W$** **Regression Equation**

$$\bar{\varepsilon} = 43.1073 - 103.359 \cdot \bar{c} + 22.3868 \cdot W + 65.389 \cdot \bar{c}^2 + 0.00117845 \cdot W^2 - 37.8818 \cdot \bar{c} \cdot W$$

**Coefficients**

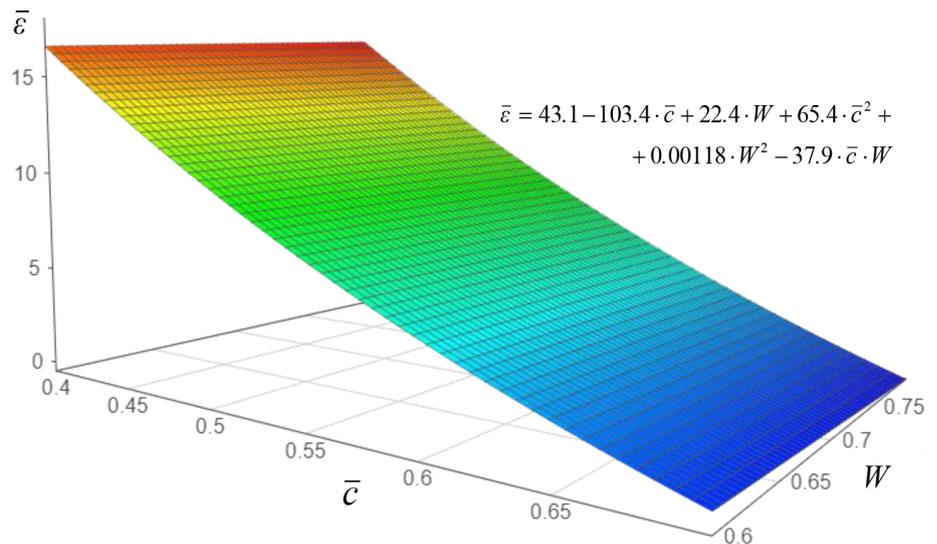
Term	Coef	SE Coef	T	P
Constant	43.107	1.50909	28.5650	0.000
$\bar{c}$	-103.359	3.35368	-30.8196	0.000
$W$	22.387	3.20286	6.9896	0.000
$\bar{c}^2$	65.389	2.15008	30.4124	0.000
$W^2$	0.001	0.00027	4.4336	0.000
$\bar{c} \cdot W$	-37.882	2.33708	-16.2091	0.000

**Summary of Model**

S = 0.415199 R-Sq = 68.86% R-Sq(adj) = 68.82%  
 PRESS = 684.233 R-Sq(pred) = 68.73%

**Analysis of Variance**

Source	DF	SS	Adj SS	Adj MS	F	P	$\eta^2$
Regression	5	320007	320007	64001.4	3261.66	0.0000000	
$\bar{c}$	1	287840	14946	14945.9	761.68	0.0000000	0.71132
$W$	1	16091	1344	1344.0	68.49	0.0000000	0.03976
$\bar{c}^2$	1	9614	11366	11366.0	579.24	0.0000000	0.02376
$W^2$	1	276	339	338.9	17.27	0.0000331	0.00068
$\bar{c} \cdot W$	1	6185	6185	6185.3	315.22	0.0000000	0.01528
Error	4314	84651	84651	19.6			
Lack-of-Fit	3833	82686	82686	21.6	5.28	0.0000000	
Pure Error	481	1964	1964	4.1			
Total	4319	404658					

**FIGURE 14** Graphical representation of the experimental regressive model in Equation (16)

The final regression equation is

$$\bar{\varepsilon} \cong 43.1 - 103.4 \cdot \bar{c} + 22.4 \cdot W + 65.4 \cdot \bar{c}^2 + 0.00118 \cdot W^2 - 37.9 \cdot \bar{c} \cdot W. \quad (16)$$

The residual plots in Figure 12 show that the variance of the residuals seems rather uniform across the full range of fitted values, confirming the efficacy of the above “weighing” of the observations. Although the top-right plot denotes a slight under-fitting pattern in the bottom-left part, residuals seem globally satisfactory. Furthermore, they can be considered as randomly distributed by the Anderson–Darling normality test at  $p < 0.05$ .

The regression output is quantitatively examined by an ANOVA (see Figure 13). Based on a  $t$  test at  $p < 0.05$ , it can be deduced that

all the terms in Equation (16) are significant. The model fits the experimental data well.

In addition to the significance tests, the so-called *effect size* of the terms in Equation (16) can be estimated (Levine & Hullett, 2002). Precisely, the ANOVA table in Figure 13 can be enriched by determining the *eta-squared* coefficient related to each term:

$$\eta^2 = \frac{SS_{term}}{SS_{Total}} \quad (17)$$

SS and  $SS_{Total}$  values are reported in the third column of the ANOVA table itself. From a practical point of view,  $\eta^2$  describes the proportion of variance of the dependent variable ( $\bar{\varepsilon}$ ), which is explained by the term

of interest; according to a rule of thumb,  $0 \leq \eta^2 \leq 0.01$  denotes a *small* effect,  $0.01 < \eta^2 \leq 0.06$  denotes a *medium* effect, while  $0.06 < \eta^2 \leq 1$  denotes a *large* effect (Field, 2013; Pierce, Block, & Aguinis, 2004). Therefore, returning to the analysis in Figure 13, the  $\eta^2$  values reported in the lower right part denote a large effect of the  $\bar{c}$  and  $W$  terms, a medium effect of the  $\bar{c}^2$  and  $\bar{c} \cdot W$  terms, and a small effect of the  $W^2$  term.

Figure 14 graphically represents the final regression equation, confirming the previous results: the predominant effect of  $\bar{c}$ , the relatively lower effect of  $W$  on  $\bar{e}$ , and the relatively weak interaction between  $\bar{c}$  and  $W$ .

We underline that this regression model is aimed at confirming and quantifying the effects of  $\bar{c}$  and  $W$  on  $\bar{e}$ , which have already been highlighted in the “Results” section. Unfortunately, this model cannot be used for predictive purposes (e.g., to estimate the  $\bar{e}$  value of a certain incomplete problem), for the simple fact that  $W$  cannot be calculated directly for incomplete problems, but only for complete ones.

## 5 | CONCLUDING REMARKS

The  $ZM_{II}$ -technique includes a flexible response mode that can be adapted to various practical contexts, in which (a) the concentration of judges cannot realistically be too high, or (b) judges may find it difficult to formulate complete preference rankings. This flexibility encourages the reliability of input data, as it prevents judges from providing forced and unreliable responses. The present study proposed an original approach to verify the robustness of the  $ZM_{II}$ -technique through massively structured experimentation.

Based on a large number of simulated decision-making problems (i.e., 4,320), the proposed factorial experiments showed the robustness of the aggregation technique, even for significantly incomplete problems. The concept of *robustness* is interpreted here as the ability to provide a relatively stable solution, despite relatively significant variations in the type of input data. Interestingly, the technique tends to converge to a reasonable solution, even for problems where judges can identify only a few more preferred objects, neglecting the others (e.g., Type- $t$  rankings with  $t = 1$  or 2). The solution towards which the technique converges is the one related to complete rankings; the plausibility of the  $ZM_{II}$ 's solution is guaranteed by the fact that it coincides with that of the LCJ, which is a very consolidated technique of the scientific literature. For a quantitative assessment of the plausibility of the results related to a specific complete problem – without necessarily knowing the corresponding complete problem from which it derives – it is possible to integrate the  $ZM_{II}$ -technique with some indicators present in the literature, such as those proposed in (Franceschini & Maisano, 2015) or others (Brasil Filho, Pinheiro, Coelho, & Costa, 2009; Kendall, 1962; Perny, 1998).

The factor that mostly affects the solution accuracy ( $\bar{e}$ ) is  $\bar{c}$ , depicting the degree of completeness of a problem, while the predominant sub-factor affecting  $\bar{c}$  is the “Ranking type”: in fact, the solution accuracy and dispersion are significantly deteriorated in the presence of Type- $t$  rankings. In addition, both the sub-factors, “ $t/b$  value” and

“Order  $t/b$ -block?,” have a certain impact on  $\bar{c}$ , while the sub-factor “Manage  $o_Z/o_M$ ,” depicting the ability to include  $o_Z$  and  $o_M$  within preference rankings, does not seem to be very influential.

The factor  $W$ , depicting the inter-judge agreement, has a less pronounced influence on  $\bar{e}$  than  $\bar{c}$ . Paradoxically, a relatively low degree of agreement tends to improve the solution, both in terms of accuracy and precision; however, this effect is rather weak.

The simulated (incomplete) problems are characterized by “homogeneous” preference rankings, that is, all rankings have the same form of incompleteness (e.g., all Type- $t$  rankings with  $t = 2$ , unordered  $t$ -objects and without  $o_Z/o_M$ ). Nevertheless, the aggregation technique can also be applied to problems characterized by “heterogeneous” preference rankings (e.g., partly complete, partly incomplete, and/or with different forms of incompleteness). The example in the section “Application of the aggregation technique to “heterogeneous” preference rankings” (in the Appendix A) demonstrates the adaptability of the aggregation technique to problems with heterogeneous preference rankings.

The solutions of the simulated decision-making problems were determined using an ad hoc software application, developed in the MS Excel – Visual Basic for Applications environment, which is available on request. This software application made the generation and solution of the (thousands of) simulated problems agile.

A methodological limitation of the proposed study regards the estimation of the degree of inter-judge agreement: since  $W$  is applicable to complete problems only, it was assumed that one incomplete problem “inherits” the  $W$  value of the corresponding (source) complete problem. Regarding the future, we plan to replace  $W$  with another suitable indicator, which can also be applied to incomplete problems.

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## ENDNOTES

<sup>1</sup> It is because of these two dummy objects that the technique is called “ $ZM_{II}$ ” (Franceschini & Maisano, 2019a).

<sup>2</sup> This concept will be formalized and clarified below.

<sup>3</sup> The concept of *compatibility* will be clarified below.

<sup>4</sup>  $n_{reg} = n - 2$ ,  $n$  is the total number of (regular and dummy) objects.

<sup>5</sup> In this case, “ties” are represented by indifference relationships.

<sup>6</sup> Since these incomplete rankings have been obtained by deteriorating complete rankings, there may be practical cases in which it is not possible to identify  $t$ -and-only- $t$   $t$ -objects and/or  $b$ -and-only- $b$   $b$ -objects, due to indifference relationships (“ $\sim$ ”) between some of them. For example, considering the complete ranking  $o_M > (o_1 \sim o_2) > o_3 > o_4 > \dots$  and having set  $t = 1$ , it is not possible to identify one-and-only-one  $t$ -object, since  $o_1$  and  $o_2$  are tied. To overcome this ambiguity, we have conventionally opted to include all the (possibly) tied objects within the  $t$ -objects, resulting in the following incomplete ranking  $o_M > (o_1 \sim o_2) > \{o_3|o_4|\dots\}$  and actually switching

from  $t = 1$  to  $t = 2$ . With simple adaptations, the same reasoning can be extended to  $b$ -objects.

- <sup>7</sup> The points in the plot are the means of a response variable at various levels of each factor; for each level of the examined factor, the mean is calculated by averaging all the responses obtained changing the remaining factor. A reference line is drawn at the grand mean of the response data. This kind of plot is useful for comparing magnitudes of main effects (Box, Hunter, & Hunter, 1978).
- <sup>8</sup> Interaction between two factors is present when the response at a factor level depends upon the level(s) of the other factor. The greater the departure of the curves from the parallel state, the higher the degree of interaction (Box et al., 1978).
- <sup>9</sup> This coefficient is a measure of the linear correlation between two variables and has a value between +1 and -1, where +1 is total positive correlation, 0 is no correlation, and -1 is total negative correlation (Ross, 2014).
- <sup>10</sup> For example, for a given  $t/b$  value, the Type- $t$  or Type  $t&b$  rankings with a relatively large number of regular objects ( $n_{reg}$ ) are likely to have lower  $\bar{c}$  and  $W$  values. Also, when increasing the number of judges ( $m$ ), the possibility to generate discordant preference rankings will grow (and therefore  $W$  will decrease).
- <sup>11</sup> Section 5 showed that  $\sigma^2$  is positively correlated with  $\bar{e}$ . Since the variance of  $\bar{e}$  is related to  $\bar{\sigma}$  (demonstration is left to the reader), it can be deduced that the variance of  $\bar{e}$  tends to grow while  $\bar{e}$  itself increases.
- <sup>12</sup> Although being aware that  $\sigma^2$  depends on the variances of the  $\hat{y}_i$  values [see Equation (5)], we have - for simplicity - considered  $\sigma^2$  as a proxy for the point variance related to each  $\bar{e}$  value.

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APPENDIX A

A.1 Example of generation of incomplete rankings

Table A1 shows an example of deterioration of a single complete ranking – that is,  $(o_M \sim o_{12}) \succ (o_9 \sim o_8) \succ o_1 \succ (o_{11} \sim o_2) \succ (o_{10} \sim o_3 \sim o_6) \succ o_7 \succ (o_4 \sim o_5 \sim o_2)$  – into the aforementioned 19 incomplete rankings, by changing the deterioration sub-factors in Figure 4b. For each of these incomplete rankings, the corresponding  $c_k$  value is also determined to depict the degree of completeness (see the last column of Table A1); of course, the  $c_k$  values may change depending on the combination of deterioration sub-factors.

A.2 Application of the aggregation technique to “heterogeneous” preference rankings

This section exemplifies a decision-making problem in which preference rankings of different nature are aggregated (e.g., partly complete, partly incomplete, and/or with different forms of incompleteness). Let us consider the example in Table A2, which contains 20 not-necessarily-complete preference rankings (representing an incomplete

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**TABLE A1** Example of generation of 19 different incomplete preference rankings (in the last column) by “deteriorating” a single complete ranking –  $(o_M \sim o_{12}) \succ (o_9 \sim o_8) \succ o_1 \succ (o_{11} \sim o_2) \succ (o_{10} \sim o_3 \sim o_6) \succ o_7 \succ (o_4 \sim o_5 \sim o_2)$  – through different combinations of the four sub-factors: “Ranking type,” “Manage  $o_Z/o_M$ ,” “t/b value,” and “Order t/b objects?”

Ranking type	Manage $o_Z/o_M$ ?	t/b value	Order t/b-objects?	Incomplete rankings	c
Quasi-complete	No	N/A	N/A	$\{o_M  o_{12}\} \succ (o_9 \sim o_8) \succ o_1 \succ (o_{11} \sim o_2) \succ (o_{10} \sim o_3 \sim o_6) \succ o_7 \succ \{o_Z  (o_4 \sim o_5)\}$	96.7%
Type-t&b	Yes	1	Yes	$(o_M \sim o_{12}) \succ \{o_1  o_2  o_3  o_6  o_7  o_8  o_9  o_{10}  o_{11}\} \succ (o_4 \sim o_5 \sim o_2)$	60.4%
Type-t&b	Yes	2	Yes	$(o_M \sim o_{12}) \succ (o_9 \sim o_8) \succ \{o_1  o_2  o_3  o_6  o_7  o_{10}  o_{11}\} \succ (o_4 \sim o_5 \sim o_2)$	76.9%
Type-t&b	Yes	3	Yes	$(o_M \sim o_{12}) \succ (o_9 \sim o_8) \succ \{o_1  o_2  o_3  o_6  o_{10}  o_{11}\} \succ o_7 \succ (o_4 \sim o_5 \sim o_2)$	83.5%
Type-t&b	No	1	Yes	$\{o_M  o_{12}\} \succ \{o_1  o_2  o_3  o_6  o_7  o_8  o_9  o_{10}  o_{11}\} \succ \{o_Z  (o_4 \sim o_5)\}$	57.1%
Type-t&b	No	2	Yes	$\{o_M  o_{12}\} \succ (o_9 \sim o_8) \succ \{o_1  o_2  o_3  o_6  o_7  o_{10}  o_{11}\} \succ \{o_Z  (o_4 \sim o_5)\}$	73.6%
Type-t&b	No	3	Yes	$\{o_M  o_{12}\} \succ (o_9 \sim o_8) \succ \{o_1  o_2  o_3  o_6  o_{10}  o_{11}\} \succ o_7 \succ \{o_Z  (o_4 \sim o_5)\}$	80.2%
Type-t&b	N/A	1	No	$\{o_M  o_{12}\} \succ \{o_1  o_2  o_3  o_6  o_7  o_8  o_9  o_{10}  o_{11}\} \succ \{o_Z  o_4  o_5\}$	56.0%
Type-t&b	N/A	2	No	$\{o_M  o_8  o_9  o_{12}\} \succ \{o_1  o_2  o_3  o_6  o_7  o_{10}  o_{11}\} \succ \{o_Z  o_4  o_5\}$	67.0%
Type-t&b	N/A	3	No	$\{o_M  o_8  o_9  o_{12}\} \succ \{o_1  o_2  o_3  o_6  o_{10}  o_{11}\} \succ \{o_Z  o_4  o_5  o_7\}$	70.3%
Type-t	Yes	1	Yes	$(o_M \sim o_{12}) \succ \{o_Z  o_1  o_2  o_3  o_4  o_5  o_6  o_7  o_8  o_9  o_{10}  o_{11}\}$	27.5%
Type-t	Yes	2	Yes	$(o_M \sim o_{12}) \succ (o_9 \sim o_8) \succ \{o_Z  o_1  o_2  o_3  o_4  o_5  o_6  o_7  o_{10}  o_{11}\}^*$	50.5%
Type-t	Yes	3	Yes	$(o_M \sim o_{12}) \succ (o_9 \sim o_8) \succ \{o_Z  o_1  o_2  o_3  o_4  o_5  o_6  o_7  o_{10}  o_{11}\}$	50.5%
Type-t	No	1	Yes	$\{o_M  o_{12}\} \succ \{o_Z  o_1  o_2  o_3  o_4  o_5  o_6  o_7  o_8  o_9  o_{10}  o_{11}\}$	26.4%
Type-t	No	2	Yes	$\{o_M  o_{12}\} \succ (o_9 \sim o_8) \succ \{o_Z  o_1  o_2  o_3  o_4  o_5  o_6  o_7  o_{10}  o_{11}\}^*$	49.5%
Type-t	No	3	Yes	$\{o_M  o_{12}\} \succ (o_9 \sim o_8) \succ \{o_Z  o_1  o_2  o_3  o_4  o_5  o_6  o_7  o_{10}  o_{11}\}$	49.5%
Type-t	N/A	1	No	$\{o_M  o_{12}\} \succ \{o_Z  o_1  o_2  o_3  o_4  o_5  o_6  o_7  o_8  o_9  o_{10}  o_{11}\}$	26.4%
Type-t	N/A	2	No	$\{o_M  o_8  o_9  o_{12}\} \succ \{o_Z  o_1  o_2  o_3  o_4  o_5  o_6  o_7  o_{10}  o_{11}\}^*$	44.0%
Type-t	N/A	3	No	$\{o_M  o_8  o_9  o_{12}\} \succ \{o_Z  o_1  o_2  o_3  o_4  o_5  o_6  o_7  o_{10}  o_{11}\}$	44.0%

Note: Reconstructed parts are marked in red colour.  
 \*Consistently with the convention described in footnote 6, the number of t- or b-objects in these rankings is higher than the respective t and b values, due to the presence of some indifference relationships among these objects.

**TABLE A2** Incomplete preference rankings (in the last column) that are obtained by deteriorating some complete source rankings (in the second column)

		Deterioration parameters			
Judge	Complete rankings (Complete problem)	Ranking type	Manage $o_z/o_M$ ?	$t/b$ value	Order $t/b$ -block(s)?
$J_1$	$(o_M \sim o_1 \sim o_8 \sim o_{12}) \succ o_9 \succ o_7 \succ o_6 \succ o_{11} \succ o_2 \succ o_4 \succ (o_3 \sim o_{10} \sim o_z \sim o_5)$	Type-t&b	Yes	3	Yes
$J_2$	$(o_M \sim o_8) \succ o_{11} \succ o_9 \succ (o_7 \sim o_{12}) \succ o_1 \succ (o_4 \sim o_2) \succ (o_{10} \sim o_6 \sim o_3) \succ (o_5 \sim o_z)$	Type-t&b	Yes	3	No
$J_3$	$(o_M \sim o_9) \succ o_1 \succ o_{11} \succ o_{12} \succ o_8 \succ o_5 \succ (o_3 \sim o_2) \succ o_7 \succ (o_6 \sim o_{10} \sim o_4 \sim o_2)$	Type-t	No	3	Yes
$J_4$	$(o_M \sim o_9 \sim o_{12}) \succ o_8 \succ o_1 \succ (o_4 \sim o_2) \succ o_7 \succ o_{11} \succ (o_6 \sim o_3) \succ (o_{10} \sim o_z \sim o_5)$	Complete	Yes	N/A	N/A
$J_5$	$o_M \succ o_1 \succ o_9 \succ (o_8 \sim o_4) \succ (o_6 \sim o_{11} \sim o_{12}) \succ (o_2 \sim o_7) \succ o_3 \succ (o_{10} \sim o_z \sim o_5)$	Type-t&b	No	1	Yes
$J_6$	$(o_M \sim o_{12}) \succ (o_9 \sim o_8) \succ o_1 \succ (o_{11} \sim o_2) \succ (o_{10} \sim o_3 \sim o_6) \succ o_7 \succ (o_4 \sim o_5 \sim o_2)$	Type-t	Yes	3	Yes
$J_7$	$o_M \succ (o_1 \sim o_8) \succ o_9 \succ o_{12} \succ o_2 \succ (o_7 \sim o_4) \succ o_{11} \succ o_5 \succ (o_6 \sim o_{10} \sim o_z \sim o_3)$	Type-t&b	No	1	No
$J_8$	$(o_M \sim o_9) \succ (o_2 \sim o_7 \sim o_1) \succ o_5 \succ (o_8 \sim o_{12}) \succ o_{11} \succ (o_6 \sim o_4) \succ (o_{10} \sim o_z \sim o_3)$	Type-t	Yes	3	Yes
$J_9$	$(o_M \sim o_{11} \sim o_{12}) \succ o_9 \succ (o_1 \sim o_6 \sim o_7) \succ o_5 \succ (o_2 \sim o_3) \succ o_4 \succ o_8 \succ (o_z \sim o_{10})$	Type-t	Yes	3	Yes
$J_{10}$	$(o_M \sim o_{12}) \succ (o_{11} \sim o_5) \succ o_9 \succ o_7 \succ (o_4 \sim o_2) \succ (o_8 \sim o_6) \succ o_1 \succ (o_{10} \sim o_z \sim o_3)$	Type-t&b	No	1	No
$J_{11}$	$(o_M \sim o_{12}) \succ (o_7 \sim o_8 \sim o_1) \succ o_{11} \succ o_9 \succ o_2 \succ o_3 \succ (o_{10} \sim o_6 \sim o_4) \succ (o_5 \sim o_2)$	Type-t	Yes	3	Yes
$J_{12}$	$(o_M \sim o_8 \sim o_9 \sim o_{12}) \succ (o_{11} \sim o_2) \succ (o_1 \sim o_5) \succ o_6 \succ o_7 \succ (o_3 \sim o_{10}) \succ (o_z \sim o_4)$	Type-t	Yes	3	Yes
$J_{13}$	$o_M \succ o_8 \succ (o_1 \sim o_{12}) \succ o_9 \succ o_2 \succ (o_{11} \sim o_2) \succ o_7 \succ o_{10} \succ (o_z \sim o_6 \sim o_3 \sim o_4)$	Type-t&b	Yes	3	No
$J_{14}$	$(o_M \sim o_1) \succ o_8 \succ o_{12} \succ (o_2 \sim o_5) \succ (o_9 \sim o_6 \sim o_7) \succ (o_{11} \sim o_4) \succ o_3 \succ o_7 \succ (o_z \sim o_{10})$	Type-t	No	3	No
$J_{15}$	$o_M \succ o_{12} \succ (o_9 \sim o_2) \succ o_1 \succ (o_8 \sim o_6 \sim o_7) \succ (o_{11} \sim o_{10}) \succ o_5 \succ (o_4 \sim o_2 \sim o_3)$	Type-t&b	No	3	No
$J_{16}$	$(o_M \sim o_8 \sim o_9) \succ o_{12} \succ o_7 \succ (o_4 \sim o_5) \succ o_6 \succ o_1 \succ o_{11} \succ o_3 \succ o_{10} \succ o_2 \succ o_z$	Type-t	No	3	Yes
$J_{17}$	$o_M \succ o_{11} \succ o_2 \succ (o_4 \sim o_8 \sim o_1) \succ (o_5 \sim o_{12}) \succ (o_3 \sim o_9) \succ o_7 \succ o_{10} \succ (o_z \sim o_6)$	Type-t	No	2	Yes
$J_{18}$	$(o_M \sim o_9 \sim o_{12}) \succ o_1 \succ o_2 \succ (o_6 \sim o_8) \succ o_7 \succ (o_{11} \sim o_5) \succ o_4 \succ (o_{10} \sim o_z \sim o_3)$	Quasi-complete	No	N/A	N/A
$J_{19}$	$(o_M \sim o_8 \sim o_{12}) \succ o_2 \succ o_7 \succ (o_{11} \sim o_1) \succ o_9 \succ (o_6 \sim o_4) \succ o_3 \succ (o_{10} \sim o_z \sim o_5)$	Type-t	No	3	Yes
$J_{20}$	$(o_M \sim o_9 \sim o_{11}) \succ o_{12} \succ o_8 \succ o_1 \succ o_5 \succ (o_2 \sim o_7) \succ o_4 \succ (o_6 \sim o_{10} \sim o_z \sim o_3)$	Type-t&b	No	4	No

Note: It can be noticed that the deterioration mechanism may differ from ranking to ranking. The reconstructed parts of the incomplete rankings are marked in red colour.

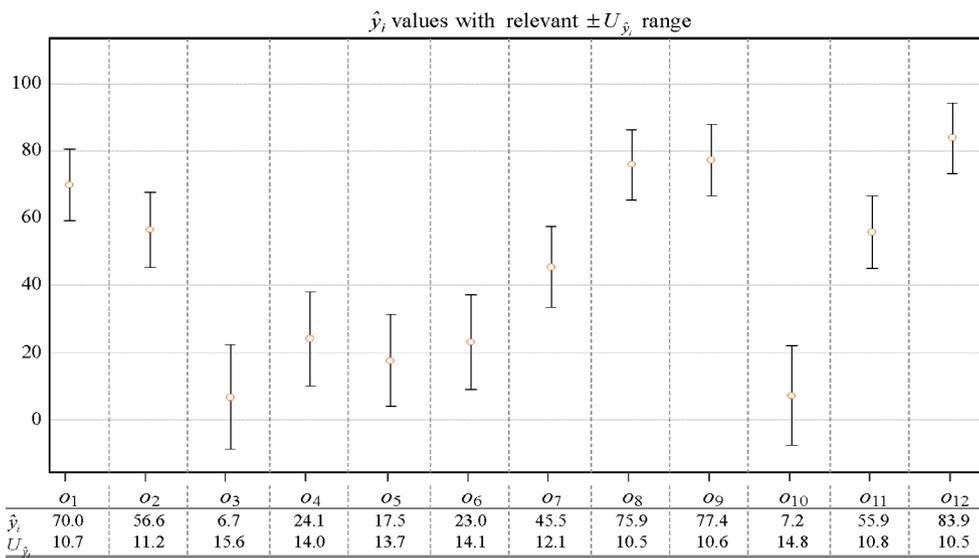
problem), obtained by deteriorating 20 corresponding complete rankings (representing the “source” complete problem). Deterioration is performed by changing the four sub-factors in Figure 4.

The aggregation technique of interest is applied to the incomplete problem, resulting in the solution in Figure A1. We note that the error bands of  $\hat{y}_i$  values (which represent  $\hat{y}_i \pm U_{\hat{y}_i}$ ; see Equation (11)) are relatively wide, due to the relatively low degree of completeness of the problem ( $\bar{c} = 62.2\%$ ). Despite this, the above solution is relatively close to that of the complete problem (which – by the way – is also characterized by  $W = 69.9\%$ ). Figure A2 shows that  $\hat{y}_i$  values related to these two solutions are strongly correlated.

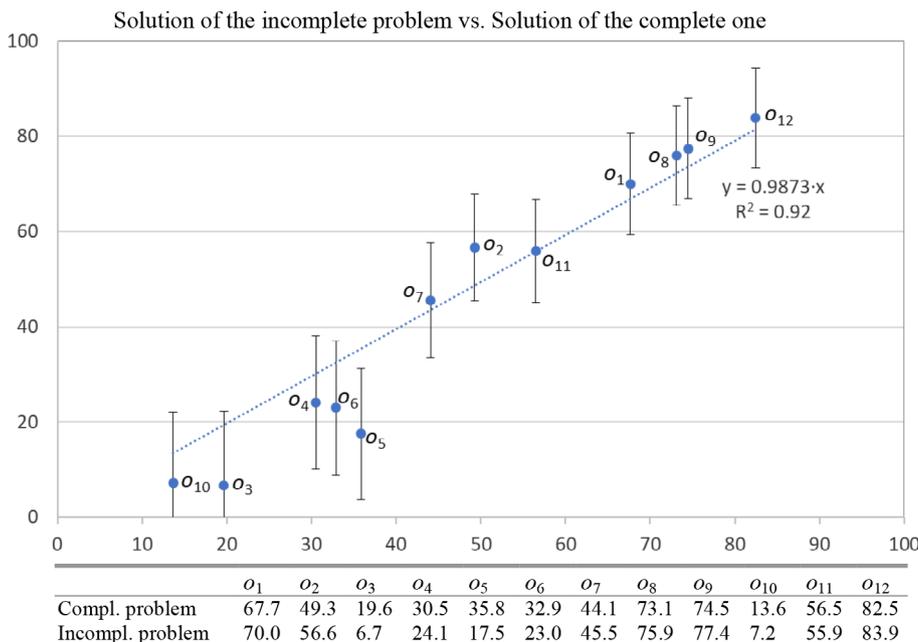
Curiously enough, the value of  $\bar{e}$  for the incomplete problem (i.e., 7.98) is relatively close to the value that would be expected through the application of the experimental regression model in Equation (16):

$$\begin{aligned} \bar{e}(\bar{c} = 62.2\%, W = 69.9\%) &= 43.1 - 103.4 \cdot (62.2\%) + 22.4 \cdot (69.9\%) \\ &+ 65.4 \cdot (62.2\%)^2 + 0.00118 \cdot (69.9\%)^2 - 37.9 \cdot (62.2\%) \cdot (69.9\%) \\ &\cong 3.27 \end{aligned} \tag{A1}$$

The above result represents a sort of plausibility check of the model in Equation (16).



**FIGURE A1** Graphical representation of the solution of the incomplete problem in Table A2 (last column)



**FIGURE A2** Comparison between the solution of the incomplete problem and that of the complete problem in Table A2. Error bands represent the expanded uncertainty associated with the scaling of the incomplete problem. The complete problem is characterized by  $W = 69.9\%$ , while the incomplete problem by  $\bar{c} = 62.2\%$  and  $\bar{e} = 7.98$